# CONSTRUCTING THE NON-LINEAR REGRESSION EQUATION TO ESTIMATE THE SOFTWARE SIZE OF OPEN SOURCE PHP-BASED INFORMATION SYSTEMS 

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## INTRODUCTION

PHP is a free programming language used primarily in information systems and web applications. Software size is one of the most important internal metrics of software including software of PHP-based open-source information systems. The information obtained from estimating the software size is useful for predicting the software development effort by such models as COCOMO 81, COCOMO II and COCOMO 2000. The papers [1, 2] proposed the linear regression equations for estimating the software size of some programming languages, such as VBA, PHP, Java and C++. The proposed equations are constructed by multiple linear regression analysis on the basis of the metrics that can be measured from class diagram. However, there are four basic assumptions that justify the use of linear regression models, one of which is normality of the error distribution. But this assumption is valid only in particular cases. This leads to the need to use the non-linear regression equa-
tions including for estimating the software size of PHPbased open-source information systems.

A normalizing transformation is often a good way to build the equations, confidence and prediction intervals of multiply non-linear regressions [3-8]. According to [4] transformations are used for essentially four purposes, two of which are: first, to obtain approximate normality for the distribution of the error term (residuals), second, to transform the response and/or the predictor in such a way that the strength of the linear relationship between new variables (normalized variables) is better than the linear relationship between dependent and independent random variables. Well-known techniques for building the equations, confidence and prediction intervals of multivariate non-linear regressions are based on the univariate normalizing transformations, which do not take into account the correlation between random variables in the case of normalization of multivariate non-Gaussian data. This leads to
the need to use the multivariate normalizing transformations.

The goal of the article is to construct the non-linear regression equation for estimating the software size of opensource PHP-based information systems. The software size prediction results by constructed equation should be better in comparison with other regression equations, both linear and nonlinear, primarily on such standard evaluations as the multiple coefficient of determination and mean magnitude of relative error.

In this article, we build the equation, confidence and prediction intervals of multivariate non-linear regression for estimating the software size of open-source PHP-based systems on the basis of the Johnson multivariate normalizing transformation (the Johnson normalizing translation) with the help of appropriate techniques proposed in $[8,9]$.

The techniques. The techniques to build the equations, confidence and prediction intervals of non-linear regressions are based on the multiple non-linear regression analysis using the multivariate normalizing transformations. A multivariate normalizing transformation of nonGaussian random vector $\mathbf{P}=\left\{\mathrm{Y}, \mathrm{X}_{1}, \mathrm{X}_{2}, \ldots, \mathrm{X}_{\mathrm{k}}\right\}^{\top}$ to Gaussian random vector $\mathbf{T}=\left\{\mathrm{Z}_{\mathrm{Y}}, \mathrm{Z}_{1}, \mathrm{Z}_{2}, \ldots, \mathrm{Z}_{\mathrm{k}}\right\}^{\top \mathrm{T}}$ is given by

$$
\begin{equation*}
\mathbf{T}=\boldsymbol{\psi}(\mathbf{P}) \tag{1}
\end{equation*}
$$

and the inverse transformation for (1)

$$
\begin{equation*}
\mathbf{P}=\boldsymbol{\psi}^{-1}(\mathbf{T}) . \tag{2}
\end{equation*}
$$

The linear regression equation for normalized data according to (1) will have the form [4]

$$
\begin{equation*}
\hat{Z}_{Y}=\bar{Z}_{Y}+\left(\mathbf{Z}_{X}^{+}\right) \hat{\mathbf{b}}, \tag{3}
\end{equation*}
$$

where $\hat{Z}_{Y}$ is prediction linear regression equation result for values of components of vector $\mathbf{z}_{\mathrm{X}}=\left\{\mathrm{Z}_{1}, \mathrm{Z}_{2}, \ldots, \mathrm{Z}_{\mathrm{k}}\right\} ; \mathbf{Z}_{\mathrm{X}}^{+}$is the matrix of centered regressors that contains the values $Z_{1_{i}}-\bar{Z}_{1}, Z_{2_{\mathrm{i}}}-\bar{Z}_{2}$,
$\ldots, Z_{k_{i}}-\bar{Z}_{k} ; \hat{\mathbf{b}}$ is estimator for vector of linear regression equation parameters, $\mathbf{b}=\left\{\mathrm{b}_{1}, \mathrm{~b}_{2}, \ldots, \mathrm{~b}_{\mathrm{k}}\right\}^{\mathrm{T}}$.

The non-linear regression equation will have the form

$$
\begin{equation*}
\hat{\mathrm{Y}}=\psi_{\mathrm{Y}}^{-1}\left[\overline{\mathrm{Z}}_{\mathrm{Y}}+\left(\mathbf{Z}_{\mathrm{X}}^{+}\right) \hat{\mathbf{b}}\right] . \tag{4}
\end{equation*}
$$

where $\psi_{\mathrm{Y}}$ is the first component of vector $\boldsymbol{\psi}=\left\{\psi_{\mathrm{Y}}, \psi_{1}, \psi_{2}, \ldots, \psi_{\mathrm{k}}\right\}^{\mathrm{T}}$.

The technique to build a confidence interval of nonlinear regression is based on transformations (1) and (2), equation (3) and a confidence interval of linear regression for normalized data

$$
\begin{equation*}
\hat{\mathbf{Z}}_{\mathrm{Y}} \pm \mathrm{t}_{\alpha / 2, \mathrm{v}} \mathrm{~S}_{\mathrm{Z}_{\mathrm{Y}}}\left\{\frac{1}{\mathrm{~N}}+\left(\mathbf{z}_{\mathrm{X}}^{+}\right)^{\mathrm{T}}\left[\left(\mathbf{Z}_{\mathrm{X}}^{+}\right)^{\mathrm{T}} \mathbf{Z}_{\mathrm{X}}^{+}\right]^{-1}\left(\mathbf{z}_{\mathrm{X}}^{+}\right)\right\}^{1 / 2} \tag{5}
\end{equation*}
$$

where $t_{\alpha / 2, v}$ is a quantile of student's $t$-distribution with $v$ degrees of freedom and $\alpha / 2$ significance level; $\left(z_{x}^{+}\right)^{T}$ is one of the rows of $\mathbf{Z}_{X}^{+} ; S_{Z_{Y}}^{2}=\frac{1}{v} \sum_{i=1}^{N}\left(Z_{Y_{i}}-\hat{Z}_{Y_{i}}\right)^{2}$, $v=\mathrm{N}-\mathrm{k}-1 ;\left(\mathbf{Z}_{\mathrm{X}}^{+}\right)^{\mathrm{T}} \mathbf{Z}_{\mathrm{X}}^{+}$is the $\mathrm{k} \times \mathrm{k}$ matrix

$$
\left(\mathbf{z}_{\mathrm{X}}^{+}\right)^{\mathrm{T}} \mathbf{z}_{\mathrm{X}}^{+}=\left(\begin{array}{cccc}
\mathrm{s}_{\mathrm{Z}_{1} \mathrm{z}_{1}} & \mathrm{~s}_{\mathrm{Z}_{1} \mathrm{z}_{2}} & \cdots & \mathrm{~s}_{\mathrm{z}_{1} \mathrm{z}_{\mathrm{k}}} \\
\mathrm{z}_{\mathrm{z}_{1} \mathrm{z}_{2}} & \mathrm{~S}_{\mathrm{z}_{2} \mathrm{z}_{2}} & \cdots & \mathrm{~s}_{\mathrm{z}_{2} \mathrm{z}_{\mathrm{k}}} \\
\cdots & \cdots & \cdots & \cdots \\
\mathrm{~s}_{\mathrm{z}_{\mathrm{k}} \mathrm{z}_{\mathrm{k}}} & \mathrm{~s}_{\mathrm{z}_{2} \mathrm{z}_{\mathrm{k}}} & \cdots & \mathrm{~s}_{\mathrm{z}_{\mathrm{k}} \mathrm{z}_{\mathrm{k}}}
\end{array}\right),
$$

where $\mathrm{S}_{\mathrm{Z}_{\mathrm{q}} \mathrm{Z}_{\mathrm{r}}}=\sum_{\mathrm{i}=1}^{\mathrm{N}}\left[\mathrm{Z}_{\mathrm{q}_{\mathrm{i}}}-\bar{Z}_{\mathrm{q}}\left[\mathrm{Z}_{\mathrm{r}_{\mathrm{i}}}-\bar{Z}_{\mathrm{r}}\right], \mathrm{q}, \mathrm{r}=1,2, \ldots, \mathrm{k}\right.$.
The confidence interval for non-linear regression is built on the basis of the interval (5) and inverse transformation (2)

$$
\begin{equation*}
\psi_{Y}^{-1}\left(\hat{Z}_{Y} \pm \mathrm{t}_{\alpha / 2, \mathrm{v}} \mathrm{~S}_{\mathrm{Z}_{\mathrm{Y}}}\left\{\frac{1}{\mathrm{~N}}+\left(\mathbf{z}_{\mathrm{X}}^{+}\right)^{\mathrm{T}}\left[\left(\mathbf{Z}_{\mathrm{X}}^{+}\right)^{\mathrm{T}} \mathbf{Z}_{\mathrm{X}}^{+}\right]^{-1}\left(\mathbf{z}_{\mathrm{X}}^{+}\right)\right\}^{1 / 2}\right) . \tag{6}
\end{equation*}
$$

The technique to build a prediction interval is based on multivariate transformation (1), the inverse transformation


$$
\begin{equation*}
\hat{Z}_{\mathrm{Y}} \pm \mathrm{t}_{\alpha / 2, \mathrm{v}} \mathrm{~S}_{\mathrm{Z}_{\mathrm{Y}}}\left\{1+\frac{1}{\mathrm{~N}}+\left(\mathbf{z}_{\mathrm{X}}^{+}\right)^{\mathrm{T}}\left[\left(\mathbf{Z}_{\mathrm{X}}^{+}\right)^{\mathrm{T}} \mathbf{Z}_{\mathrm{X}}^{+}\right]^{-1}\left(\mathbf{z}_{\mathrm{X}}^{+}\right)\right\}^{1 / 2} \tag{7}
\end{equation*}
$$

The prediction interval for non-linear regression is built on the basis of the interval (7) and inverse transformation (2)

$$
\begin{equation*}
\psi_{\mathrm{Y}}^{-1}\left(\hat{\mathrm{Z}}_{\mathrm{Y}} \pm \mathrm{t}_{\alpha / 2, \mathrm{v}} \mathrm{~S}_{\mathrm{Z}_{\mathrm{Y}}}\left\{1+\frac{1}{\mathrm{~N}}+\left(\mathbf{z}_{\mathrm{X}}^{+}\right)^{\mathrm{T}}\left[\left(\mathbf{Z}_{\mathrm{X}}^{+}\right)^{\mathrm{T}} \mathbf{Z}_{\mathrm{X}}^{+}\right]^{-1}\left(\mathbf{z}_{\mathrm{X}}^{+}\right)\right\}^{1 / 2}\right) \tag{8}
\end{equation*}
$$

The Johnson normalizing translation. For normalizing the multivariate non-Gaussian data, we use the Johnson translation system. In our case the Johnson normalizing translation is given by [10]

$$
\begin{equation*}
\mathbf{T}=\boldsymbol{\gamma}+\boldsymbol{\eta} \mathbf{h}\left[\lambda^{-1}(\mathbf{P}-\varphi)\right] \sim \mathrm{N}_{\mathrm{m}}\left(\mathbf{0}_{\mathrm{m}}, \boldsymbol{\Sigma}\right) \tag{9}
\end{equation*}
$$

where $\boldsymbol{\Sigma}$ is the covariance matrix; $\mathrm{m}=\mathrm{k}+1 ; \boldsymbol{\gamma}, \boldsymbol{\eta}, \varphi$ and $\lambda$ are parameters of translation (9); $\boldsymbol{\gamma}=\left(\gamma_{\mathrm{Y}}, \gamma_{1}, \gamma_{2}, \ldots, \gamma_{\mathrm{k}}\right)^{\mathrm{T}} ; \quad \boldsymbol{\eta}=\operatorname{diag}\left(\eta_{\mathrm{Y}}, \eta_{1}, \eta_{2}, \ldots, \eta_{\mathrm{k}}\right)$; $\boldsymbol{\lambda}=\operatorname{diag}\left(\lambda_{\mathrm{Y}}, \lambda_{1}, \lambda_{2}, \ldots, \lambda_{\mathrm{k}}\right) ; \quad \varphi=\left(\varphi_{\mathrm{Y}}, \varphi_{1}, \varphi_{2}, \ldots, \varphi_{\mathrm{k}}\right)^{\mathrm{T}}$; $\mathbf{h}\left[\left(\mathrm{y}_{\mathrm{Y}}, \mathrm{y}_{1}, \ldots, \mathrm{y}_{\mathrm{k}}\right)\right]=\left\{\mathrm{h}_{\mathrm{Y}}\left(\mathrm{y}_{\mathrm{Y}}\right), \mathrm{h}_{1}\left(\mathrm{y}_{1}\right), \ldots, \mathrm{h}_{\mathrm{k}}\left(\mathrm{y}_{\mathrm{k}}\right)\right\}^{\mathrm{T}} ; \mathrm{h}_{\mathrm{i}}(\cdot)$ is one of the translation functions

$$
\mathrm{h}=\left\{\begin{array}{cl}
\ln (\mathrm{y}), & \text { for } \mathrm{S}_{\mathrm{L}} \text { (log normal)family }  \tag{10}\\
\ln [\mathrm{y} /(1-\mathrm{y})], & \text { for } \mathrm{S}_{\mathrm{B}} \text { (bounded)family } \\
\operatorname{Arsh}(\mathrm{y}), & \text { for } \mathrm{S}_{\mathrm{U}} \text { (unbounded)family } \\
\mathrm{y} & {\text { for } \mathrm{S}_{\mathrm{N}} \text { (normal)family }}^{\text {(nor }}
\end{array}\right.
$$

Here $y=(x-\varphi) / \lambda ; \quad \operatorname{Arsh}(y)=\ln \left(y+\sqrt{y^{2}+1}\right) . \ln$ our case $X$ equals $Y, X_{1}, X_{2}$ or $X_{3}$ respectively.

The equation, confidence and prediction intervals of non-linear regression to estimate the software size of open-source PHP-based systems. The equation, confidence and prediction intervals of non-linear regression to estimate the software size of open-source PHP-based systems are constructed on the basis of the

Johnson multivariate normalizing transformation for the four-dimensional non-Gaussian data set: actual software size in the thousand lines of code ( KLOC ) Y, the average number of attributes per class $\mathrm{X}_{3}$, the total number of classes $\mathrm{X}_{1}$ and the total number of relationships $\mathrm{X}_{2}$ in conceptual data model from 32 information systems developed using the PHP programming language with HTML and SQL. Table I contains the data from [1] on four metrics of software for 32 open-source PHP-based systems.

For detecting the outliers in the data from Table 1 we use the technique based on multivariate normalizing transformations and the squared Mahalanobis distance [11]. There are no outliers in the data from Table I for 0.005 significance level and the Johnson multivariate transformation (9) for $\mathrm{S}_{\mathrm{B}}$ family. The same result was obtained in [12] for the transformation (9) for $S_{U}$ family. In [1] it was also assumed that the data contains no outliers. Although note that without using normalization, the data of system 11 is multivariate outlier, since for this data row the squared Mahalanobis distance equals to 15.44 is greater than the value of the quantile of the Chi-square distribution, which equals to 14.86 for 0.005 significance level.

Parameters of the multivariate transformation (9) for $S_{B}$ family were estimated by the maximum likelihood method. Estimators for parameters of the transformation (9) are: $\hat{\gamma}_{\mathrm{Y}}=9.63091, \hat{\gamma}_{1}=15.5355, \hat{\gamma}_{2}=25.4294$, $\hat{\gamma}_{3}=0.72801, \quad \hat{\eta}_{\mathrm{Y}}=1.05243, \quad \hat{\eta}_{1}=1.58306$, $\hat{\eta}_{2}=2.54714, \quad \hat{\eta}_{3}=0.54312, \quad \hat{\varphi}_{Y}=-1.4568$, $\hat{\varphi}_{1}=-1,8884, \quad \hat{\varphi}_{2}=-6,9746, \quad \hat{\varphi}_{3}=3.2925$, $\hat{\lambda}_{\mathrm{Y}}=153102.605, \hat{\lambda}_{1}=243051.0, \hat{\lambda}_{2}=311229.5$ and $\hat{\lambda}_{3}=13.900$. The sample covariance matrix $\mathrm{S}_{\mathrm{N}}$ of the $\mathbf{T}$ is used as the approximate moment-matching estimator of $\Sigma$

$$
\mathrm{S}_{\mathrm{N}}=\left(\begin{array}{llll}
1.0000 & 0.9514 & 0.9333 & 0.1574 \\
0.9514 & 1.0000 & 0.9006 & 0.1345 \\
0.9333 & 0.9006 & 1.0000 & 0.0554 \\
0.1574 & 0.1345 & 0.0554 & 1.0000
\end{array}\right)
$$

After normalizing the non-Gaussian data by the multivariate transformation (9) for $S_{B}$ family the linear regression equation (3) is built for normalized data


$$
\begin{equation*}
\hat{\mathrm{Z}}_{\mathrm{Y}}=\hat{\mathrm{b}}_{0}+\hat{\mathrm{b}}_{1} \mathrm{Z}_{1}+\hat{\mathrm{b}}_{2} \mathrm{Z}_{2}+\hat{\mathrm{b}}_{3} \mathrm{Z}_{3} . \tag{11}
\end{equation*}
$$

Parameters of the linear regression equation (11) were estimated by the least square method. Estimators for pa-
rameters of the equation (11) are such: $\hat{b}_{0}=1.02 \cdot 10^{-5}$, $\hat{\mathrm{b}}_{1}=0.56085, \hat{\mathrm{~b}}_{2}=0.42491, \hat{\mathrm{~b}}_{3}=0.05846$.

Table I

## The data and prediction result by regression equations for 32 open-source PHP-based systems

| No | Y | $\mathrm{X}_{1}$ | $\mathrm{X}_{2}$ | $\mathrm{X}_{3}$ | Linear regression |  | Non-linear regression |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  |  |  |  |  | Y | RME | univariate transformation |  | multivariate transformation |  |
|  |  |  |  |  |  |  | $\hat{\text { Y }}$ | RME | $\hat{Y}$ | RME |
| 1 | 3.038 | 5 | 2 | 10.6 | 3.237 | 0.0656 | 4.675 | 0.5388 | 4.550 | 0.4976 |
| 2 | 22.599 | 17 | 7 | 7 | 24.142 | 0.0683 | 19.965 | 0.1166 | 19.990 | 0.1154 |
| 3 | 32.243 | 21 | 13 | 4.524 | 37.524 | 0.1638 | 32.098 | 0.0045 | 33.535 | 0.0401 |
| 4 | 16.164 | 13 | 11 | 7.077 | 25.916 | 0.6033 | 23.171 | 0.4335 | 21.292 | 0.3173 |
| 5 | 83.862 | 35 | 24 | 6.571 | 74.624 | 0.1102 | 80.265 | 0.0429 | 83.618 | 0.0029 |
| 6 | 24.22 | 13 | 9 | 8.077 | 23.224 | 0.0411 | 20.524 | 0.1526 | 18.901 | 0.2196 |
| 7 | 63.929 | 35 | 19 | 8.029 | 67.215 | 0.0514 | 65.913 | 0.0310 | 70.647 | 0.1051 |
| 8 | 2.543 | 5 | 3 | 9.4 | 4.127 | 0.6228 | 5.789 | 1.2764 | 5.169 | 1.0328 |
| 9 | 6.697 | 5 | 5 | 7 | 5.906 | 0.1181 | 7.353 | 0.0980 | 6.356 | 0.0509 |
| 10 | 55.537 | 25 | 14 | 8.64 | 46.843 | 0.1565 | 42.098 | 0.2420 | 43.126 | 0.2235 |
| 11 | 55.752 | 39 | 10 | 9.077 | 57.814 | 0.0370 | 67.070 | 0.2030 | 49.823 | 0.1064 |
| 12 | 62.602 | 30 | 17 | 7 | 56.995 | 0.0896 | 53.497 | 0.1454 | 56.651 | 0.0951 |
| 13 | 67.111 | 23 | 22 | 14.957 | 61.856 | 0.0783 | 65.500 | 0.0240 | 60.617 | 0.0968 |
| 14 | 2.552 | 3 | 1 | 8.333 | -2.395 | 1.9384 | 2.202 | 0.1370 | 2.447 | 0.0412 |
| 15 | 12.17 | 10 | 5 | 3.7 | 9.959 | 0.1816 | 9.693 | 0.2035 | 10.029 | 0.1759 |
| 16 | 12.757 | 13 | 9 | 5 | 21.218 | 0.6632 | 18.682 | 0.4644 | 18.105 | 0.4192 |
| 17 | 5.695 | 7 | 3 | 8.429 | 5.976 | 0.0493 | 7.083 | 0.2438 | 6.687 | 0.1743 |
| 18 | 7.744 | 9 | 6 | 9.222 | 13.991 | 0.8067 | 12.911 | 0.6673 | 11.301 | 0.4593 |
| 19 | 7.514 | 4 | 1 | 8 | -1.371 | 1.1825 | 2.496 | 0.6678 | 3.096 | 0.5880 |
| 20 | 11.054 | 9 | 9 | 3.667 | 15.385 | 0.3918 | 13.301 | 0.2032 | 12.850 | 0.1625 |
| 21 | 29.77 | 17 | 15 | 3.412 | 35.179 | 0.1817 | 27.321 | 0.0823 | 29.061 | 0.0238 |
| 22 | 11.653 | 9 | 8 | 8.778 | 17.045 | 0.4627 | 15.461 | 0.3268 | 13.268 | 0.1386 |
| 23 | 6.847 | 5 | 4 | 3.6 | 2.017 | 0.7054 | 5.435 | 0.2062 | 5.112 | 0.2534 |
| 24 | 13.389 | 7 | 5 | 11.714 | 11.462 | 0.1440 | 10.367 | 0.2257 | 8.661 | 0.3531 |
| 25 | 14.45 | 12 | 6 | 16.583 | 22.513 | 0.5580 | 20.191 | 0.3973 | 15.888 | 0.0995 |
| 26 | 4.414 | 6 | 3 | 3.667 | 1.630 | 0.6307 | 5.318 | 0.2048 | 5.260 | 0.1916 |
| 27 | 2.102 | 3 | 1 | 3.333 | -5.655 | 3.6902 | 2.142 | 0.0192 | 1.873 | 0.1090 |
| 28 | 42.819 | 20 | 18 | 3.5 | 43.975 | 0.0270 | 37.967 | 0.1133 | 38.631 | 0.0978 |
| 29 | 4.077 | 4 | 2 | 9 | 0.953 | 0.7662 | 3.892 | 0.0454 | 3.732 | 0.0846 |
| 30 | 57.408 | 33 | 14 | 9.242 | 57.164 | 0.0043 | 53.121 | 0.0747 | 54.381 | 0.0527 |
| 31 | 7.428 | 7 | 3 | 7 | 5.044 | 0.3209 | 6.861 | 0.0764 | 6.571 | 0.1154 |
| 32 | 8.947 | 15 | 5 | 4 | 16.360 | 0.8285 | 12.934 | 0.4456 | 14.258 | 0.5936 |

After that the non-linear regression equation (4) is built where $\hat{\mathrm{Z}}_{\mathrm{Y}}$ is prediction result by the equation (11),

$$
\begin{equation*}
\hat{\mathrm{Y}}=\hat{\varphi}_{\mathrm{Y}}+\hat{\lambda}_{\mathrm{Y}}\left[1+\mathrm{e}^{-\left(\hat{z}_{\mathrm{Y}}-\hat{\gamma}_{\mathrm{Y}}\right) / \hat{\eta}_{\mathrm{Y}}}\right]^{-1} \tag{12}
\end{equation*}
$$

$Z_{j}=\gamma_{j}+\eta_{j} \ln \frac{X_{j}-\varphi_{j}}{\varphi_{j}+\lambda_{j}-X_{j}}, \quad \varphi_{j}<X_{j}<\varphi_{j}+\lambda_{j}$,
$j=1,2,3$.

The prediction results by equation (12) for values of components of vector $\mathbf{X}=\left\{\mathrm{X}_{1}, \mathrm{X}_{2}, \mathrm{X}_{3}\right\}$ from Table I and values of magnitude of relative error MRE are shown in the Table I for two cases: the Johnson univariate and multivariate normalizing transformations. Table I also contains the prediction results by linear regression equation from [1] for values of components of vector $\mathbf{X}$ from Table I and MRE values. Note the prediction results by linear regression equation from [1] are negative for the three rows of data: 14, 19 and 27. All prediction results by non-linear regression equation (12) are positive.

For univariate normalizing transformations (10) of $\mathrm{S}_{B}$ family the estimators for parameters are such: $\hat{\gamma}_{\mathrm{Y}}=0.77502, \quad \hat{\gamma}_{1}=0.59473, \quad \hat{\gamma}_{2}=0.57140$, $\hat{\gamma}_{3}=0.68734, \quad \hat{\eta}_{Y}=0.44395, \quad \hat{\eta}_{1}=0.48171$,
$\hat{\eta}_{2}=0.49553, \quad \hat{\eta}_{3}=0.51970, \quad \hat{\varphi}_{Y}=2.063$,
$\hat{\varphi}_{1}=2.900, \hat{\varphi}_{2}=0.900, \hat{\varphi}_{3}=3.304, \hat{\lambda}_{\mathrm{Y}}=83.059$, $\hat{\lambda}_{1}=36.695, \hat{\lambda}_{2}=23.525$ and $\hat{\lambda}_{3}=13.660$. In the case of univariate normalizing transformations the estimators for parameters of the equation (11) are such: $\hat{b}_{0}=3.11 \cdot 10^{-7}, \quad \hat{b}_{1}=0.43519, \quad \hat{b}_{2}=0.52239$ and $\hat{b}_{3}=0.08546$.

Also the non-linear regression equation (4) is built by the decimal logarithm transformation

$$
\begin{equation*}
\hat{Y}=10^{\mathrm{b}_{0}} X_{1}^{\mathrm{b}_{1}} X_{2}^{\mathrm{b}_{2}} X_{3}^{\mathrm{b}_{3}}, \tag{13}
\end{equation*}
$$

where the estimators for parameters of the equation (13) are such: $\hat{b}_{0}=-0.26161, \hat{b}_{1}=0.99151$, $\hat{b}_{2}=0.33232$ and $\hat{b}_{3}=0.13777$.

The values of multiple coefficient of determination $\mathrm{R}^{2}$, mean magnitude of relative error MMRE and percentage of prediction $\operatorname{PRED}(0.25)$ equal respectively $0.9491,0.4919$ and 0.5313 for linear regression equation from [1], and equal respectively $0.9375,0.2455$ and 0.625 for the equation (13). The values of $\mathrm{R}^{2}$, MMRE and $\operatorname{PRED}(0.25)$ are better for the equation (12), in comparison with both the equation from [1] and equation (13), and are $0.9692,0.2199$ and 0.7188 for the Johnson multivariate transformation and equal $0.9591,0.2535$ and 0.7188 for the Johnson univariate transformation respectively. The acceptable values of MMRE and
$\operatorname{PRED}(0.25)$ are not more than 0.25 and not less than 0.75 respectively. The values of MMRE indicate that only the values for equation (12) on the basis of the Johnson multivariate normalizing transformation and the decimal logarithm transformation are less than 0.25 . Although all values of $\operatorname{PRED}(0.25)$ are less than 0.75 nevertheless the values are greater for equation (12) with estimators of parameters for the Johnson transformations, both multivariate and univariate.

The confidence and prediction intervals of non-linear regression are defined by (6) and (8) respectively for the data from Table I. Table II contains the lower (LB) and upper (UB) bounds of the confidence intervals of linear and non-linear regressions on the basis of univariate and multivariate transformations respectively for 0.05 significance level. Note the lower bounds of the confidence interval of linear regression from [1] are negative for the seven rows of data: $1,14,19,23,26,27$ and 29 . The upper bound for the data row 27 is negative too. All the lower and upper bounds of the confidence interval of non-linear regressions are positive. The widths of the confidence interval of non-linear regression on the basis of the Johnson multivariate transformation are less than for linear regression from [1] for the twenty rows of data: $1,6,8,9,14-20,22-27,29,31$ and 32 . Also the widths of the confidence interval of non-linear regression on the basis of the Johnson multivariate transformation are less for more data rows than for non-linear regressions following the univariate transformations, both decimal logarithm and the Johnson. The widths of the confidence interval of non-linear regression on the basis of the Johnson multivariate transformation are less than following the decimal logarithm univariate transformation for the twenty-seven rows of data: 1-4, 6-12, 15-26, 28, 30-32. And ones are less than following the Johnson univariate transformation for the twenty-five rows of data: 1-4, 6, $8-11,15-18,20-26,28-32$. Approximately the same results are obtained for the prediction intervals of regressions.

Table III contains the lower (LB) and upper (UB) bounds of the prediction intervals of linear and nonlinear regressions on the basis of univariate and multivariate transformations respectively for 0.05 significance level. Note the lower bounds of the prediction interval of linear regression from [1] are negative for the thirteen

| - |  |  |
| :---: | :---: | :---: |
| ::\%:\%: | \# 23 (2018) |  |
| - |  |  |

rows of data: $1,8,9,14,15,17,19,23,24,26,27,29,31$. All the lower bounds of the prediction interval of nonlinear regressions are positive. The widths of the prediction interval of non-linear regression on the basis of the Johnson multivariate transformation are less than for linear regression from [1] for the twenty rows of data: 1, $6,8,9,14-20,22-27,29,31$ and 32 . Also the widths of the prediction interval of non-linear regression on the basis of the Johnson multivariate transformation are less for more data rows than for non-linear regressions fol-
lowing the univariate transformations, both decimal logarithm and the Johnson. The widths of the prediction interval of non-linear regression on the basis of the Johnson multivariate transformation are less than following the decimal logarithm univariate transformation for the twenty-nine rows of data: 1-13, 15-18, 20-26, 28-32. And ones are less than following the Johnson univariate transformation for the twenty-three rows of data: 1-4, 6, $8-10,15-18,20-26,28,29,31$ and 32.

Table II
Bounds of the confidence intervals

| No | Y | Bounds for linear regression |  | Bounds for non-linear regressions |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  |  |  |  | univariate transformations |  |  |  | Johnson multivariate transformation |  |
|  |  |  |  | decimal logarithm |  | Johnson |  |  |  |
|  |  | LB | UB | LB | UB | LB | UB | LB | UB |
| 1 | 3.038 | -0.402 | 6.877 | 3.725 | 5.947 | 3.673 | 6.267 | 3.655 | 5.601 |
| 2 | 22.599 | 21.413 | 26.871 | 18.933 | 27.172 | 15.473 | 25.455 | 16.856 | 23.660 |
| 3 | 32.243 | 34.344 | 40.704 | 26.415 | 39.621 | 24.791 | 40.266 | 29.157 | 38.539 |
| 4 | 16.164 | 23.172 | 28.660 | 16.855 | 24.285 | 17.982 | 29.365 | 18.542 | 24.421 |
| 5 | 83.862 | 69.187 | 80.062 | 55.076 | 87.173 | 74.107 | 83.078 | 68.095 | 102.603 |
| 6 | 24.22 | 21.015 | 25.433 | 16.557 | 22.438 | 16.129 | 25.819 | 16.935 | 21.075 |
| 7 | 63.929 | 62.690 | 71.740 | 52.309 | 83.045 | 56.434 | 72.961 | 59.182 | 84.277 |
| 8 | 2.543 | 1.013 | 7.241 | 4.288 | 6.544 | 4.484 | 7.748 | 4.260 | 6.223 |
| 9 | 6.697 | 3.084 | 8.728 | 4.569 | 7.951 | 5.456 | 10.203 | 5.135 | 7.803 |
| 10 | 55.537 | 43.863 | 49.824 | 35.573 | 52.195 | 33.947 | 50.366 | 37.344 | 49.768 |
| 11 | 55.752 | 49.560 | 66.068 | 41.448 | 87.698 | 42.891 | 79.359 | 36.261 | 68.257 |
| 12 | 62.602 | 53.265 | 60.725 | 43.512 | 65.766 | 44.275 | 61.787 | 48.298 | 66.405 |
| 13 | 67.111 | 54.146 | 69.566 | 35.747 | 69.156 | 49.897 | 75.572 | 45.429 | 80.723 |
| 14 | 2.552 | -5.673 | 0.883 | 1.639 | 2.897 | 2.125 | 2.375 | 1.806 | 3.214 |
| 15 | 12.17 | 6.609 | 13.309 | 8.838 | 13.632 | 6.979 | 13.684 | 8.320 | 12.037 |
| 16 | 12.757 | 18.574 | 23.862 | 15.339 | 21.222 | 14.673 | 23.576 | 16.253 | 20.151 |
| 17 | 5.695 | 3.165 | 8.787 | 6.139 | 8.644 | 5.548 | 9.233 | 5.655 | 7.870 |
| 18 | 7.744 | 11.381 | 16.601 | 10.039 | 14.139 | 9.902 | 16.849 | 9.960 | 12.799 |
| 19 | 7.514 | -4.587 | 1.845 | 2.106 | 3.945 | 2.253 | 3.046 | 2.391 | 3.930 |
| 20 | 11.054 | 11.684 | 19.085 | 9.137 | 15.776 | 9.186 | 19.255 | 10.550 | 15.591 |
| 21 | 29.77 | 30.767 | 39.591 | 20.246 | 34.593 | 16.796 | 41.072 | 22.899 | 36.782 |
| 22 | 11.653 | 14.250 | 19.840 | 10.430 | 16.253 | 11.581 | 20.525 | 11.347 | 15.478 |
| 23 | 6.847 | -1.579 | 5.613 | 3.900 | 6.687 | 4.071 | 7.662 | 4.063 | 6.360 |
| 24 | 13.389 | 7.648 | 15.276 | 7.099 | 11.493 | 7.462 | 14.583 | 7.262 | 10.286 |
| 25 | 14.45 | 16.199 | 28.828 | 13.006 | 22.695 | 10.971 | 34.746 | 12.197 | 20.576 |
| 26 | 4.414 | -1.967 | 5.227 | 4.421 | 7.029 | 4.083 | 7.261 | 4.241 | 6.461 |
| 27 | 2.102 | -9.730 | -1.580 | 1.360 | 2.712 | 2.092 | 2.281 | 1.087 | 2.902 |
| 28 | 42.819 | 38.873 | 49.077 | 25.212 | 43.588 | 25.181 | 51.940 | 30.491 | 48.845 |
| 29 | 4.077 | -2.236 | 4.142 | 2.947 | 4.616 | 3.177 | 5.048 | 2.959 | 4.640 |
| 30 | 57.408 | 52.335 | 61.993 | 44.400 | 73.879 | 41.599 | 63.278 | 44.841 | 65.885 |
| 31 | 7.428 | 2.314 | 7.774 | 6.042 | 8.344 | 5.463 | 8.784 | 5.585 | 7.695 |
| 32 | 8.947 | 12.515 | 20.205 | 12.503 | 22.001 | 9.080 | 18.449 | 11.119 | 18.180 |

Table III
Bounds of the prediction intervals

| No | Y | Bounds for linear regression |  | Bounds for non-linear regressions |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  |  |  |  | univariate transformations |  |  |  | Johnson multivariate transformation |  |
|  |  |  |  | decimal logarithm |  | Johnson |  |  |  |
|  |  | LB | UB | LB | UB | LB | UB | LB | UB |
| 1 | 3.038 | -8.886 | 15.361 | 2.264 | 9.787 | 2.507 | 15.664 | 2.053 | 8.822 |
| 2 | 22.599 | 12.260 | 36.024 | 11.075 | 46.451 | 5.800 | 53.204 | 11.088 | 35.207 |
| 3 | 32.243 | 25.530 | 49.517 | 15.704 | 66.644 | 9.341 | 65.987 | 19.149 | 57.962 |
| 4 | 16.164 | 14.031 | 37.802 | 9.874 | 41.455 | 6.642 | 57.342 | 11.955 | 37.129 |
| 5 | 83.862 | 61.845 | 87.403 | 33.367 | 143.886 | 59.920 | 84.392 | 47.603 | 146.045 |
| 6 | 24.22 | 11.451 | 34.998 | 9.475 | 39.211 | 5.956 | 53.906 | 10.617 | 32.866 |
| 7 | 63.929 | 54.797 | 79.633 | 31.724 | 136.931 | 31.210 | 81.247 | 40.528 | 122.355 |
| 8 | 2.543 | -7.849 | 16.103 | 2.565 | 10.940 | 2.713 | 20.215 | 2.431 | 9.838 |
| 9 | 6.697 | -5.998 | 17.810 | 2.856 | 12.722 | 2.996 | 26.099 | 3.097 | 11.949 |
| 10 | 55.537 | 34.901 | 58.785 | 20.980 | 88.499 | 13.397 | 72.304 | 24.761 | 74.346 |
| 11 | 55.752 | 43.606 | 72.022 | 27.405 | 132.638 | 26.251 | 82.571 | 26.759 | 91.726 |
| 12 | 62.602 | 44.844 | 69.146 | 25.940 | 110.319 | 19.861 | 77.358 | 32.563 | 97.782 |
| 13 | 67.111 | 47.957 | 75.755 | 23.063 | 107.190 | 28.542 | 81.562 | 33.153 | 109.857 |
| 14 | 2.552 | -14.415 | 9.625 | 1.030 | 4.613 | 2.084 | 2.994 | 0.811 | 5.262 |
| 15 | 12.17 | -2.080 | 21.999 | 5.307 | 22.704 | 3.441 | 33.425 | 5.255 | 18.197 |
| 16 | 12.757 | 9.355 | 33.081 | 8.849 | 36.788 | 5.492 | 51.258 | 10.150 | 31.513 |
| 17 | 5.695 | -5.925 | 17.877 | 3.565 | 14.884 | 2.964 | 24.822 | 3.336 | 12.381 |
| 18 | 7.744 | 2.136 | 25.846 | 5.831 | 24.343 | 4.145 | 40.894 | 6.095 | 20.095 |
| 19 | 7.514 | -13.374 | 10.632 | 1.346 | 6.172 | 2.127 | 4.916 | 1.198 | 6.351 |
| 20 | 11.054 | 3.243 | 27.527 | 5.697 | 25.303 | 4.154 | 42.480 | 6.867 | 23.133 |
| 21 | 29.77 | 22.801 | 47.556 | 12.581 | 55.670 | 7.324 | 63.400 | 15.978 | 51.960 |
| 22 | 11.653 | 5.148 | 28.943 | 6.285 | 26.973 | 4.693 | 46.152 | 7.200 | 23.590 |
| 23 | 6.847 | -10.093 | 14.128 | 2.426 | 10.749 | 2.635 | 19.103 | 2.367 | 9.829 |
| 24 | 13.389 | -0.715 | 23.638 | 4.334 | 18.826 | 3.576 | 35.238 | 4.477 | 15.796 |
| 25 | 14.45 | 9.337 | 35.689 | 8.136 | 36.280 | 5.323 | 56.560 | 8.396 | 29.076 |
| 26 | 4.414 | -10.481 | 13.741 | 2.683 | 11.585 | 2.621 | 18.450 | 2.464 | 10.048 |
| 27 | 2.102 | -17.916 | 6.606 | 0.885 | 4.168 | 2.073 | 2.648 | 0.410 | 4.484 |
| 28 | 42.819 | 31.335 | 56.615 | 15.725 | 69.883 | 10.895 | 70.978 | 21.432 | 68.748 |
| 29 | 4.077 | -11.043 | 12.949 | 1.779 | 7.647 | 2.371 | 12.014 | 1.575 | 7.423 |
| 30 | 57.408 | 44.632 | 69.696 | 27.354 | 119.916 | 19.170 | 77.441 | 30.902 | 94.883 |
| 31 | 7.428 | -6.838 | 16.926 | 3.483 | 14.475 | 2.926 | 23.959 | 3.273 | 12.168 |
| 32 | 8.947 | 4.173 | 28.547 | 7.842 | 35.078 | 4.090 | 41.560 | 7.530 | 26.021 |

Following [13] multivariate kurtosis $\beta_{2}$ is estimated for the data on metrics of software from Table I and the normalized data on the basis of the decimal logarithm transformation, the Johnson univariate and multivariate transformations for $\mathrm{S}_{\mathrm{B}}$ family. The estimator of multivariate kurtosis given by [13]

$$
\begin{equation*}
\hat{\beta}_{2}=\frac{1}{\mathrm{~N}} \sum_{\mathrm{i}-1}^{\mathrm{N}}\left\{\left(\mathbf{Z}_{\mathrm{i}}-\overline{\mathbf{Z}}\right)^{\mathrm{T}} \mathrm{~S}_{\mathrm{N}}^{-1}\left(\mathbf{Z}_{\mathrm{i}}-\overline{\mathbf{Z}}\right)\right\}^{2} . \tag{14}
\end{equation*}
$$

In our case, in the formula (14), the vectors $\mathbf{Z}$ and $\overline{\mathbf{Z}}$ should be replaced by the vectors $\mathbf{P}$ and $\overline{\mathbf{P}}$ or $\mathbf{T}$ and $\overline{\mathbf{T}}$, respectively, for the initial (non-Gaussian) or normalized
data. It is known that $\beta_{2}=m(m+2)$ holds under multivariate normality. The given equality is a necessary condition for multivariate normality. In our case $\beta_{2}=24$. The estimators of multivariate kurtosis equal 28.66, 23.87, 37.29 and 23.08 for the data from Table I, the normalized data on the basis of the decimal logarithm transformation, the Johnson univariate and multivariate transformations respectively. The values of these estimators indicate that the necessary condition for multivariate normality is practically performed for the normalized data on the basis of the decimal logarithm transformation and the Johnson multivariate transformation, it does not hold for other data. Note that in our case, the poor normalization of multivariate
non-Gaussian data using the Johnson univariate transformation leads to an increase in the widths of the confidence and prediction intervals of non-linear regression for a larger number of data rows compared to both the Johnson multivariate transformation and the decimal logarithm transformation.

## CONCLUSIONS

The non-linear regression equation to estimate the software size of open-source PHP-based information systems is improved on the basis of the Johnson multivariate transformation for $\mathrm{S}_{\mathrm{B}}$ family. This equation, in comparison with other regression equations (both linear and nonlinear), has a larger multiple coefficient of determination and a smaller value of MMRE.

When building the equations, confidence and prediction intervals of non-linear regressions for multivariate non-Gaussian data to estimate the software size of opensource PHP-based information systems, one should use multivariate transformations.

Usually poor normalization of multivariate nonGaussian data or application of univariate transformations instead of multivariate ones to normalize such data may lead to increase of width of the confidence and prediction intervals of regressions, both linear and non-linear, to estimate the software size of open-source PHP-based information systems.

In the future, we intend to try other multivariate normalizing transformations and non-Gaussian data sets.

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