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## Stock prices and exchange rates in Sri Lanka: some empirical evidence

### Abstract

This paper examines the relationship between stock prices and exchange rates in Sri Lanka. The author uses monthly data on four foreign exchange rates and the All Share Price Index (ASPI) of the Colombo Stock Exchange in the empirical analysis. The Johansen's cointegration test finds no long-run relationships between stock prices (ASPI) and any of the four exchange rates during the sample period. Therefore, the paper proceeded to test for short-run in-sample causal relationships between stock prices and exchange rates and found one unidirectional relationship from stock prices to the US dollar exchange rate. In addition, the author performed a variance decomposition analysis to get insights into the out-of-sample causal relationships between exchange rates and stock prices. The results indicate that most of the variance of the ASPI, particularly, at longer horizons is explained by the Indian rupee with the other currencies explaining a very little variation of the ASPI. The study also examined whether the ASPI has any role in explaining the variations in any of the four exchange rates. It is found that the ASPI has the most explanatory power in relation to the US dollar exchange rate.

**Keywords:** Ng-Perron test, foreign exchange market, Japanese yen, Sri Lanka, Colombo Stock Exchange.

**JEL Classification:** F31, G14, G15.

### Introduction

The causal relationship between exchange rates and stock prices has received considerable attention in both developed and developing countries. This interest emanated with the growth in international trade and financial liberalization. As a result of international trade activities, firms have been exposed to foreign exchange risk. Fluctuations in exchange rates affect input prices and assets denominated in foreign currencies (Bodnar and Gentry, 1993). Therefore, share prices of firms engaged in foreign trade are affected by the changes in profits due to exchange rate fluctuations. The share prices of firms with no foreign trade may also be indirectly affected as these firms interact domestically with those firms that are engaged in foreign trade. In contrast, Bahmani-Oskooee and Sohrabian (1992) argue that changes in stock prices may affect exchange rates through firms' portfolio adjustments. Qiao (1996) also provides a similar argument that capital outflows affect exchange rates if changes in stock prices are sufficiently persistent to generate or destroy confidence of stock market investors.

There have been a number of studies examining the causal nexus between stock prices and exchange rates. These studies provided mixed results. Some of these studies are reviewed below. Abdalla and Murinde (1997) investigated the interaction between stock prices and the real effective exchange rates in India, Korea, Pakistan and the Philippines. They found unidirectional causalities from exchange rates to stock prices in all sample countries but not in the Philippines. Granger et al. (2000) studied the causal rela-

tionships between stock prices and exchange rates using the Asian flu data. They reported that the data from South Korea are consistent with the traditional approach that exchange rates lead stock prices. However, the data for the Philippines were consistent with the portfolio approach: stock prices lead exchange rates with negative correlation. Data from Hong Kong, Malaysia, Singapore, Thailand and Taiwan indicated strong feedback relations. The data for Indonesia and Japan did not find any causal relationship.

Ibrahim (2000) examined both bivariate and multivariate cointegration and Granger causality between stock prices and exchange rates in Malaysia. This study used three measures of exchange rate, real effective exchange rate, nominal effective exchange rate and the exchange rate between ringgit and the US dollar. Although this study did not find any long-run or cointegrating relationship between exchange rate measures and stock prices in the bivariate models, there was cointegration between measures of exchange rate and stock prices when money supply ( $M2$ ) and reserves were included in the cointegrating relationship. Findings of the multivariate model indicated that, in the short run, a concerted stance on monetary, exchange rate and reserve policies is vital for stock market stability. Yong and Isa (2000) also studied the relationship between exchange rates and stock prices in Malaysia with the objective of finding whether currency depreciation during the East Asian crisis affected the stock market prices. They reported that exchange rates tended to move in tandem with the stock prices during the currency crisis. This indicates that a depreciation of the Malaysian ringgit was accompanied by a decrease in stock prices. However, this result was not consistent with that for

the period prior to the currency crisis. They point out that this may have been due to the erosion of investor confidence in the Malaysian currency. They used Granger causality test as the major econometric tool in the empirical analysis. However, they have not examined the time series properties of the variables used before conducting Granger causality tests. This raises questions about the validity of their results.

Hong (2002) examined whether Malaysian stock prices respond asymmetrically to interest rates and exchange rates using cointegration and error-correction models. This study found for the models for the stock price and interest rates that positive deviations from equilibrium tend to have a greater effect than the negative deviations. The model for the exchange rate indicated that the exchange rate does not respond to own shocks, stock price and interest rate deviations from equilibrium. Baharumshah et al. (2002) applied an augmented monetary model to the ringgit/US dollar exchange rate and ringgit/yen exchange rate. They found that the augmented monetary model was cointegrated. However, this model was subject to parameter instability and parameter time dependency could be attributed at least partly to a particular subset of variables in the system including stock prices. They found that a restricted vector auto regression model which imposes exogeneity restrictions on  $I(1)$  variables, such as stock prices, among others, exhibits both cointegration and parameter stability. They showed that exchange rate adjusted to clear any disequilibrium in the long-run relationship. The findings of this study suggest that the equity market is significant in affecting exchange rate and in explaining at least in part the parameter instability evidenced in the cointegrating system. Therefore, they suggested that the models of equilibrium exchange rate should be extended to include equity markets in addition to bond markets.

Wu (2000) focussed on the asymmetric effect of four different exchange rates on Singapore stock prices and the effects' sensitivity to economic instability. The results of this study found that Singapore currency appreciation against the US dollar and Malaysian ringgit and depreciation against Japanese yen and Indonesian rupiah lead to a long-run increase in stock prices in most of the selected periods in the 1990s. However, during the 1997-98 crisis period and the 1999-2000 period, the effect associated with the US dollar displayed a sign reversal. This study further found that the influence of exchange rates on stock prices increased in a chronological order in the 1990s.

The dynamic relationship between stock prices and exchange rates in G-7 countries was examined in a study by Nieh and Lee (2001). This study did not find

any long-run significant relationship between stock prices and exchange rates. They found a short-run significant relationship only for one day in certain G-7 countries. For example, currency depreciation often dragged down stock returns in the German financial market, but it stimulated Canadian and the UK markets on the following day. However, an increase in stock price often caused currency depreciation the next day in Italy and Japan. Kim (2003) studied the long-run equilibrium relationships among stock price, industrial production, real exchange rate, interest rate and inflation in the United States. This study found that the S&P 500 index is positively related to the industrial production but negatively to the real exchange rate, interest rate and inflation. Error-correction models estimated revealed that the stock price, industrial production and inflation adjust to correct disequilibrium among the five variables. Variance decomposition analysis showed that the stock price was driven to a considerable extent by innovations in interest rate.

Phylaktis and Ravazzollo (2005) focussed on the dynamic relationship between stock prices and exchange rates in the pacific countries. This study found that stock and foreign exchange markets are positively related and that the US stock market acted as a conduit for these links. Furthermore, these links were not caused by foreign exchange restrictions. The recursive estimates provided evidence that the financial crisis had a temporary impact on the long-run co-movement of these markets.

Mishra (2004) investigated the relationships between stock and foreign exchange markets in India. This study found evidence for a unidirectional causality between exchange rate return and interest rate and between exchange rate return and demand for money. The variance decomposition analysis performed indicated that the exchange rate return affects demand for money, the interest rate causes exchange rate return change, exchange rate return affects the stock return, the demand for money affects stock return, interest rate affects the stock return and the demand for money affects interest rate.

The main objective of this paper is to examine the causal relationships between stock market prices and exchange rates in Sri Lanka. To author's knowledge, no previous study has undertaken such a task. The results of this study will have implications for both the participants of the stock and foreign exchange markets in Sri Lanka. If causal relationships are detected from stock prices to exchange rates, the participants in the foreign exchange market can devise rules to predict the movements of exchange rates. If causal relationships are found from exchange rates, participants of the stock market can devise rules to predict the movements of stock prices.

Predictability of exchange rates or stock prices may lead to profitable trading strategies. Further, the predictability of exchange rates or stock prices violates the applicability of the efficient markets hypothesis in its semi-strong form.

The rest of this paper is organized as follows. Section 1 outlines the methodology and data used, section 2 presents the analysis of empirical results and the final section concludes the paper.

## 1. Methodology and data

Data analysis in this paper is carried out in five steps. In the first step, unit root tests are performed to examine whether the selected exchange rates and share indices are stationary in the first differences or follow random walks. The second step, conducting cointegration tests among exchange rates, is conditional on the results of the first step. That is, it can be performed providing at least one of the variables is stationary in the first differences (Hansen and Juselius, 2002). Therefore, we proceed to the second step if the above condition is satisfied. According to Engle and Granger (1987), if two variables are cointegrated, the relationships among variables can be modeled using an error-correction model for the variables. The above is the third step in the analysis. In the fourth step, Granger causality tests are performed using the error-correction model estimated in step three. Then, in the fifth step, we perform a variance decomposition analysis. The above steps are briefly discussed below<sup>1</sup>.

**1.1. Ng-Perron unit root tests.** Ng and Perron (2001) constructed four unit root test statistics that are calculated using the generalized least squares (GLS) detrended data for a variable. Compared to the widely-used Dickey-Fuller (DF) and Phillips-Perron (PP) unit root tests, these ones have better power and size properties. The first unit root test statistic developed by Ng and Perron ( $MP_T^p$ ) calculates the Elliot, Rothenberg and Stock (ERS) point optimal statistic for GLS detrended data. The other three statistics,  $MZ_{\alpha}^d$ ,  $MZ_t^d$  and  $MSB^d$  are the enhancements of the PP test statistics which correct for size distortions when residuals are negatively correlated. All four test statistics above are based on a specification for  $x_t$  and a method for estimating  $f_0$ , the zero frequency spectrum term. The specification for  $x_t$  can take one of two forms. That is, a constant or a constant and a linear trend. The consistent estimate of the residual spectrum at frequency zero is obtained on the basis of autoregressive (AR) spectral regression (GLS-detrended).

<sup>1</sup> These tests have been used in many empirical papers and, therefore, they are now well-known. Hence, they are not discussed in detail in this paper to save space. In addition, more details on these techniques can be obtained from the works cited in this section.

**1.2. Johansen's multivariate cointegration test.** This paper adopts Johansen (1991, 1995) cointegration test to test for cointegration among the exchange rates and stock prices. In making inferences about the number of cointegrating relations, two statistics known as the trace statistic and the maximal eigenvalue statistic are used. To make inferences regarding the number of cointegrating relationships, the trace and maximum eigenvalue statistics are compared with the critical values tabulated in Osterwald-Lenum (1992).

**1.3. Error-correction model, short- and long-run causality.** As mentioned earlier, according to Engle and Granger (1987), if the variables are cointegrated, the relationship among them can be modelled using an error-correction model. The error-correction model opens up another channel of causality through the error-correction term which is ignored in standard Granger causality tests. Therefore, causality can also be tested by examining (1) the statistical significance of the error-correction term by a separate  $t$ -test; (2) the joint significance of the lags of each explanatory variable by an  $F$ - or Wald  $\chi^2$  test; or by testing (3) the error-correction terms and lagged term of each explanatory variable simultaneously by a joint  $F$ - or Wald  $\chi^2$  test.

**1.4. Variance decomposition analysis.** Granger causality test results can be interpreted as within-sample causality tests and can be used to make inferences about causal relationships within the sample period only. Therefore, to make inferences on causal relationships beyond the sample period, variance decomposition analysis is used. In variance decomposition analysis, variance of the forecast error of a particular variable is partitioned into proportions attributable to innovations (or shocks) in each variable in the system, including its own. If a variable can be optimally forecast from its own lags, then it will have all its forecast variance accounted for by its own disturbances (Sims, 1982).

**1.5. Data.** Data used in this study consist of All Share Price Index (ASPI) of the Colombo Stock Exchange, exchange rates for the Indian rupee (IR), the Japanese yen (JPY), the UK pound (UKP) and the US dollar (USD) expressed as the amounts of Sri Lankan rupees per unit of each currency on a monthly basis from January 1986 to December 2004. Data on the ASPI were obtained from the data bases of the Colombo Stock Exchange. Data on exchange rates were obtained from the Monthly Bulletins published by the Central Bank of Sri Lanka.

## 2. Analysis of empirical results

Table 1 presents the descriptive statistics for the natural log values of the ASPI and four exchange rates used in the study. A comparison of mean and medians for the variables indicates that median is less than

mean for the Indian rupee, the UK pound and the US dollar. However, it is the opposite for the ASPI and the Japanese yen. This indicates that the distributions of the Indian rupee, the UK pound and the US dollar are positively skewed. In other words there have been a large number of small values and a small number of large values for these exchange rates during the sample period. As far as the ASPI and the Japanese yen are concerned, their distributions are negatively skewed as their means are less than their medians. This indicates that there has been a large number of large values and a small number of small values for the ASPI and the Japanese yen during the sample period. In other words, there has been a large number of months with large stock returns and a small number of months with small stock returns. The coefficient of skewness also supports the above conclusions.

Table 1. Descriptive statistics for the variables

|                              | ASPI   | IR     | JPY     | UKP   | USD    |
|------------------------------|--------|--------|---------|-------|--------|
| Mean                         | 6.206  | 0.622  | -0.819  | 4.468 | 3.986  |
| Median                       | 6.364  | 0.580  | -0.726  | 4.402 | 3.923  |
| Maximum                      | 7.336  | 0.869  | 0.010   | 5.308 | 4.652  |
| Minimum                      | 4.837  | 0.396  | -1.984  | 3.670 | 3.314  |
| Range                        | 2.499  | 0.472  | 1.993   | 1.638 | 1.338  |
| Standard deviation           | 0.667  | 0.155  | 0.491   | 0.395 | 0.400  |
| Coefficient of variation (%) | 10.752 | 24.906 | -59.999 | 8.851 | 10.028 |

|             |                     |                     |                     |       |                     |
|-------------|---------------------|---------------------|---------------------|-------|---------------------|
| Skewness    | -0.633              | 0.197               | -0.339              | 0.046 | 0.051               |
| Kurtosis    | 2.378               | 1.377               | 2.096               | 2.351 | 1.904               |
| Jarque-Bera | 18.905 <sup>a</sup> | 26.505 <sup>a</sup> | 12.136 <sup>a</sup> | 4.084 | 11.517 <sup>a</sup> |

Notes: <sup>a</sup>Implies statistical significance at the one per cent level. ASPI, IR, JPY, UKP and USD denote All Share Price Index, the Indian rupee, the Japanese yen, the UK pound and the US dollar, respectively.

The highest maximum values are observed for the ASPI, the UK pound exchange rate and the US dollar exchange rate. These values are due to the magnitudes of the original series. For example, out of the four exchange rates, the UK pound exchange rate has the highest maximum as we need a higher amount of Sri Lankan rupees per one pound than the amount of Sri Lankan rupees needed to purchase one unit of the other currencies. This is true for the minimum values as well. The magnitude of the original natural log values is the highest for the ASPI and, therefore, it has the highest maximum, minimum and range. A perusal of the standard deviations indicates that the ASPI has the highest standard deviation and the Indian rupee/Sri Lankan rupee has the lowest standard deviation. As standard deviation is an absolute measure of variation, to compare the series we can look at the coefficient of variation, which is a relative measure of variation. Accordingly, the Japanese yen has the highest variation during the sample period followed by the Indian rupee.

Table 2. Unit root test results

| Variable      | Unit root test statistic |                     |                    |                    |
|---------------|--------------------------|---------------------|--------------------|--------------------|
|               | $MZ_a$                   | $MZ_t$              | $MS_b$             | $MPT$              |
| ASPI          | 0.905                    | 0.819               | 0.905              | 57.463             |
|               | -3.789                   | -1.376              | 0.363              | 24.045             |
| $\Delta$ ASPI | -101.306 <sup>a</sup>    | -7.115 <sup>a</sup> | 0.070 <sup>a</sup> | 0.246 <sup>a</sup> |
|               | -107.951 <sup>a</sup>    | -7.347 <sup>a</sup> | 0.068 <sup>a</sup> | 0.845 <sup>a</sup> |
| IR            | -1.559                   | -0.859              | 0.551              | 15.249             |
|               | -1.444                   | -0.564              | 0.391              | 35.880             |
| $\Delta$ IR   | -106.846 <sup>a</sup>    | -7.263 <sup>a</sup> | 0.068 <sup>a</sup> | 0.314 <sup>a</sup> |
|               | -111.996 <sup>a</sup>    | -7.470 <sup>a</sup> | 0.067 <sup>a</sup> | 0.861 <sup>a</sup> |
| JPY           | 1.259                    | 2.175               | 1.728              | 205.614            |
|               | -4.977                   | -1.548              | 0.311              | 18.162             |
| $\Delta$ JPY  | -0.196                   | -0.184              | 0.937              | 47.619             |
|               | -14.128                  | -2.634 <sup>c</sup> | 0.186 <sup>c</sup> | 6.595 <sup>c</sup> |
| UKP           | 1.838                    | 3.170               | 1.725              | 228.979            |
|               | -8.031                   | -1.949              | 0.243              | 11.513             |
| $\Delta$ UKP  | -126.391 <sup>a</sup>    | -7.921 <sup>a</sup> | 0.063 <sup>a</sup> | 0.243 <sup>a</sup> |
|               | -141.892 <sup>a</sup>    | -8.410 <sup>a</sup> | 0.059 <sup>a</sup> | 0.686 <sup>a</sup> |
| USD           | 1.561                    | 3.620               | 2.319              | 387.525            |
|               | -9.535                   | -2.180              | 0.229              | 9.572              |
| $\Delta$ USD  | -65.752 <sup>a</sup>     | -5.733 <sup>a</sup> | 0.087 <sup>a</sup> | 0.374 <sup>a</sup> |
|               | -133.462 <sup>a</sup>    | -8.165 <sup>a</sup> | 0.061 <sup>a</sup> | 0.695 <sup>a</sup> |

Notes: See notes for Table 1 for the definitions of the notations used in column 1.  $\Delta$  indicates the first difference of the variables. <sup>a,b,c</sup>Imply statistical significance at the 1%, 5% and 10% level, respectively. The first figure under each unit root test for a currency/ASPI or its first difference is the unit root test statistics when a constant is used as the deterministic component. The second figure is the unit root test statistic when a constant and a linear time trend are used as deterministic components. The lag lengths in the Ng-Perron tests were selected using spectral GLS-detrended based on SIC.

The results of the four Ng-Perron unit root tests are shown in Table 2. Column two shows the results for  $MZ_a$  test. The results for the levels of the variables show that none of the variables is stationary. However, the results for the first differences indicate that all variables except Japanese yen exchange rate are stationary. Therefore, we can conclude that except for the Japanese yen exchange rate, all the other variables have unit roots or follow random walks. Columns three to five of the table show the results for  $MZ_b$ ,  $MSB$  and  $MPT$  tests, respectively. They also show that all variables except for the Japanese yen exchange rate are random walks when both a

constant and a constant and a time trend are considered as deterministic components in the test equations. The Japanese yen exchange rate has a random walk only when a constant and a time trend are considered as deterministic components. According to Hansen and Juselius (2002), it is possible to find cointegration among the variables, if at least two variables out of all variables considered in cointegration tests are integrated of order one or have unit roots. As the unit root test results reported in Table 2 satisfy this condition, it is possible to proceed to the second step of the analysis, testing for cointegration among stock market prices and exchange rates.

Table 3. Johansen cointegration test results for exchange rates and stock prices

| Pairs of currencies | Number of lags in VAR | Trend assumption | Null hypothesis | Trace statistic | Maximal eigenvalue statistic |
|---------------------|-----------------------|------------------|-----------------|-----------------|------------------------------|
| ASPI and IR         | 2                     | 1                | $r = 0$         | 7.603           | 6.906                        |
|                     |                       |                  | $r \leq 1$      | 0.697           | 0.697                        |
| ASPI and JPY        | 5                     | 2                | $r = 0$         | 7.699           | 5.415                        |
|                     |                       |                  | $r \leq 1$      | 2.284           | 2.284                        |
| ASPI and UKP        | 8                     | 2                | $r = 0$         | 6.665           | 6.654                        |
|                     |                       |                  | $r \leq 1$      | 0.011           | 0.011                        |
| ASPI and USD        | 3                     | 2                | $r = 0$         | 9.487           | 9.370                        |
|                     |                       |                  | $r \leq 1$      | 0.117           | 0.117                        |

Notes: See note for Table 1 for definitions of notations in column 1. The deterministic components are selected using the Pantula principle suggested by Johansen (1992). In the trend assumption column, 1 denotes no deterministic trend (restricted constant), 2 denotes linear deterministic trend, and 3 denotes linear deterministic trend (restricted). Lag lengths in vector autoregressions were selected using Likelihood Ratio (LR) test. Critical values for the trace and maximal eigenvalue tests were obtained from Osterwald-Lenum (1992).

Table 3 shows the results of Johansen cointegration test for the exchange rates and stock prices. Column two of the table shows the number of lags in the vector autoregression selected using the likelihood ratio test. Column three of the table shows the trend assumption made in cointegration tests. As cointegration test results are sensitive to the trend assumption made, the trend assumption was selected using the Pantula Principle suggested by Johansen (1992). Most previous studies on stock price exchange rate relationship and in other areas have arbitrarily chosen the trend assumption in cointegration tests. This raises questions about the validity of the results of these studies. According to the results reported in columns five and six, both null hypotheses shown in column four are rejected. Therefore, we can conclude that there is no cointegration relationship between stock prices and exchange rates in Sri Lanka.

Table 4. Granger causality test results

| Causality |      | $\chi^2$ test statistic | Direction of causality |
|-----------|------|-------------------------|------------------------|
| From      | To   |                         |                        |
| ASPI      | IR   | 0.001                   | No causality           |
| IR        | ASPI | 0.584                   |                        |
| ASPI      | JPY  | 0.001                   | No causality           |
| JPY       | ASPI | 0.209                   |                        |
| ASPI      | UKP  | 11.080                  | No causality           |
| UKP       | ASPI | 6.489                   |                        |

|      |      |                    |             |
|------|------|--------------------|-------------|
| ASPI | USD  | 5.918 <sup>c</sup> | ASPI to USD |
| USD  | ASPI | 0.003              |             |

As we did not find any cointegrating relationship between stock prices and any of the exchange rates, we examined whether we could find any short-run causal relationships between stock prices and exchange rates. Table 4 shows these results. The first column of the table shows the cause variable whereas the second column shows the effect variable. The third column of the table shows the Chi-square statistic test results to test the null hypothesis that there is no causality from the variables shown in the first column to the variables shown in the second column. According to the results, the hypothesis of no causality is not rejected for ASPI and Indian rupee, ASPI and Japanese yen, and ASPI and the UK pound. The null hypothesis of no causality is rejected only for the ASPI and USD exchange rate. Accordingly, there is a unidirectional causality from the ASPI to the US dollar exchange rate. This indicates that the current US dollar exchange rate can be predicted from the past values of the ASPI.

Table 5. Percentage of forecast variance in the ASPI

| Months | Explained by innovations in |       |       |       |
|--------|-----------------------------|-------|-------|-------|
|        | IR                          | JPY   | UKP   | USD   |
| 1      | 0.000                       | 0.000 | 0.000 | 0.000 |
| 12     | 2.548                       | 0.157 | 0.104 | 0.064 |

Table 5 (cont.). Percentage of forecast variance in the ASPI

| Months | Explained by innovations in |       |       |       |
|--------|-----------------------------|-------|-------|-------|
|        | IR                          | JPY   | UKP   | USD   |
| 24     | 7.302                       | 0.449 | 0.710 | 0.139 |
| 36     | 13.239                      | 1.482 | 2.129 | 0.516 |
| 48     | 19.522                      | 2.895 | 4.388 | 1.214 |

Granger causality tests indicate only the within-sample causality. To gain insights into out-of-sample causalities between stock prices and the four exchange rates, we performed a variance decomposition analysis. In the variance decomposition analysis, variance of the forecast error of a particular variable is partitioned into proportions attributable to innovations (or shocks) in each variable in the system, including its own. Table 5 reports the percentage of forecast variance in the ASPI explained by the innovations of each currency. According to the results, most of the variance of the ASPI, particularly, at longer horizons is explained by the innovations in the Indian rupee. For example, Indian rupee explains approximately 13 per cent and 20 per cent of the variance of the ASPI at horizons thirty six and forty eight, respectively. As far as the Japanese yen and the UK pound are concerned, they explain less than 4 per cent of the variance of the ASPI. For example, Japanese yen explains approximately 3 per cent of the variance ASPI at horizon forty eight whereas the UK pound explains 4 per cent of the variance of the ASPI at the same horizon. The US dollar exchange rate explains a very low percentage of the forecast variance of the ASPI. For example, it explains approximately 1 per cent of the variance of the ASPI at horizon forty eight.

Table 6. Percentage of forecast variance in exchange rates explained by innovations in the ASPI

| Months | IR    | JPY   | UKP   | USD    |
|--------|-------|-------|-------|--------|
| 1      | 0.897 | 0.072 | 0.578 | 1.018  |
| 12     | 0.638 | 0.627 | 5.415 | 16.554 |
| 24     | 0.579 | 1.707 | 3.192 | 23.427 |
| 36     | 0.533 | 4.509 | 3.664 | 28.252 |
| 48     | 0.494 | 7.449 | 6.171 | 31.846 |

Table 6 shows the forecast variance of each exchange rate explained by the innovations in the ASPI. According to the results, the ASPI explains a very little percentage of the Indian rupee. At horizon one, the ASPI explains approximately 0.9 per cent of forecast variance of the Indian rupee whereas it explains only 0.5 per cent of the variance of the Indian rupee at horizon forty eight. As far as the Japanese yen exchange rate is concerned, the explanatory power of the ASPI goes up with the increase in the time horizon. However, the maximum amount of variance in the Japanese yen exchange rate that

the ASPI can explain is approximately seven per cent. The ASPI can explain a maximum of approximately six per cent of the forecast variance of the UK pound exchange rate. The ASPI has the highest degree of explanatory power in relation to the USD dollar exchange rate. At horizon one, the ASPI explains approximately one per cent of forecast variance of the US dollar. This amount increases to 16 per cent at horizon twelve and at horizon 48 it is approximately 31 per cent. Overall, the above results show that there is a unidirectional causality from Indian rupee to the ASPI and the ASPI to the US dollar out of sample.

## Conclusion

In this paper we investigated the relationship between stock prices and the exchange rates of the Sri Lankan rupee against four currencies, namely the Indian rupee, the Japanese yen, the UK pound and the US dollar. Unit roots of the above five variables were examined using econometric techniques which are more powerful than the widely used Dickey-Fuller type unit root tests. As we found that all variables have unit roots, we proceeded to examine whether there are any long-run relationships between stock prices and exchange rates in Sri Lanka. For this purpose, we employed Johansen's multiple cointegration test paying careful attention to the selection of lags in the cointegration tests which has been overlooked in most empirical studies in the area of the study as well as in studies in other areas. Our tests indicated that there is no long-run relationship between stock prices and exchange rates in Sri Lanka. This finding prompted us to investigate whether there are any short-run relationships between stock prices and exchange rates of Sri Lanka using Granger-causality tests. These tests found one unidirectional causal relationship from stock prices in Sri Lanka to the US dollar exchange rate.

As Granger-causality tests indicate only within-sample causal relationships, we use variance decomposition analysis to examine whether there are out-of-sample causal relations between stock prices and exchange rates. The results indicated that most of the variance of the ASPI, particularly, at longer horizons is explained by the Indian rupee and the other currencies explain a very little variation of the ASPI. We also examined whether the ASPI has any role in explaining the variations in any of the four exchange rates. We found that the ASPI has the most explanatory power in relation to the US dollar exchange rate of the Sri Lankan rupee.

The above results indicate that stock prices in Sri Lanka can be used to predict movements of the US dollar exchange rate both within and outside the sample period. Further, the Indian rupee exchange

rate can be used to predict the movements of stock prices at long horizons out-of-sample. Therefore, the participants of foreign exchange market can devise strategies to gain from foreign exchange transaction using stock prices to predict the movements of the US dollar exchange rate both within and outside the sample period. However, the participants of the stock market can devise strategies to make gains using the Indian rupee exchange rate to predict stock prices only out of the sample period. But, both the

participants in the stock and foreign exchange markets need to consider the transaction costs involved before they decide to engage in any trading activities. The above results also violate the validity of the efficient market hypothesis to the stock and foreign exchange markets in Sri Lanka. Further, the above results provide mixed evidence for the validity of the traditional approach or portfolio approach in relation to the relationship between exchange rates and stock prices in Sri Lanka.

## References

1. Abdalla, I.S.A. and Murinde (1997). Exchange Rate and Stock Price Interactions in Emerging Markets: Evidence on India, Korea, Pakistan and the Philippines, *Applied Financial Economics*, 7, pp. 25-35.
2. Baharumshah, A.Z., Mashih, A.M.M and Azali, M. (2002). The Stock Market and the Ringgit Exchange Rate: A Note, *Japan and the World Economy*, 14, pp. 471-486.
3. Bahmani-Oskooee, M. and Sohrabian, A. (1992). Stock Prices and the Effective Exchange Rate of the Dollar, *Applied Economics*, 24, pp. 459-64.
4. Bodnar, G.M. and Gentry, W.M. (1993). Exchange rate exposure and industry characteristics: Evidence from Canada, Japan and the US, *Journal of International Money and Finance*, 12, pp. 29-45.
5. Engle, R.F., and Granger, C.W.J. (1987). "Cointegration and Error Correction Representation, Estimation and Testing", *Econometrica*, 55, pp. 251-76.
6. Granger, C.J., Huang, B. and Yang, C. (2000). A Bivariate Causality between Stock Prices and Exchange Rates: Evidence from Recent Asian Flu, *Quarterly Review of Economics and Finance*, 40, pp. 337-354.
7. Hansen, H. and Juselius, K. (2002). *CATS in RATS Cointegration Analysis of Time Series*, Illinois: Estima.
8. Hong, T.J. (2002). Do Malaysian Stock Prices Respond Asymmetrically to Interest Rate and Foreign Exchange Rate Shocks? An Empirical Evidence, *Banker's Journal Malaysia*, No. 121, pp. 35-42.
9. Ibrahim, M. (2000). Cointegration and Granger Causality Tests of Stock Price and Exchange Rate Interactions in Malaysia, *ASEAN Economic Bulletin*, 17, pp. 36-47.
10. Johansen, S. (1991). "Estimation and Hypothesis Testing of Cointegration Vectors in Gaussian Vector Autoregressive Models", *Econometrica*, 59, pp. 1551-1580.
11. Johansen, S. (1995). *Likelihood-based Inference in Cointegrated Vector Autoregressive Models*, Oxford: Oxford University Press.
12. Kim, K. (2003). Dollar Exchange Rate and Stock Price: Evidence from Multivariate Cointegration and Error Correction Model, *Review of Financial Economics*, 12, pp. 301-313.
13. Mishra, A.K. (2004). Stock Market and Foreign Exchange Market in India: Are They Related? *South Asia Economic Journal*, 5, pp. 209-232.
14. Ng, S., Perron, P. (2001). "Lag Length Selection and the Construction of Unit Root Tests with Good Size and Power", *Econometrica*, 69, pp. 1519-1554.
15. Nieh, C. and Lee, C. (2001). Dynamic Relationship between Stock Prices and Exchange Rates for G-7 Countries, *Quarterly Review of Economics and Finance*, 41, pp. 477-490.
16. Osterwald-Lenum, M. (1992). "A Note With Quantiles of the Asymptotic Distribution of the Maximum Likelihood Cointegration Rank Test Statistics", *Oxford Bulletin of Economics and Statistics*, 54, pp. 461-472.
17. Phylaktis, K. and Ravazzolo, F. (2005). Stock Prices and Exchange Rate Dynamics, *Journal of International Money and Finance*, 24, pp. 1031-1053.
18. Sims, C.A. (1982). Policy Analysis with Econometric Models, *Brookings Papers on Economic Activity*, 1, pp. 107-152.
19. Wu, Y. (2000). Stock Prices and Exchange Rates in a VEC Model – The Case of Singapore in the 1990s, *Journal of Economics and Finance*, 24, pp. 260-274.
20. Yong, O.B. and Isa, Z.B. (2000). Exchange Rate Changes and Stock Market Reactions: A Special Case of Currency Depreciation and its Influence on the local stock market: The Malaysian Experience, *Banker's Journal Malaysia*, 113, pp. 56-63.