# ІНФОРМАЦІЙНІ ТЕХНОЛОГІЇ В ЕКОНОМІЦІ 

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# DETERMINATION OF THE PRODUCTION CURVES FOR THE CONSTRUCTION INDUSTRY IN THE CZECH REPUBLIC 


#### Abstract

The construction industry is the first to be affected by developments in the economy. It is the first to contract during a period of economic recession and the first to expand during a period of economic growth. This brings to mind whether it is possible to predict the development of the economy based on predictions for the construction industry. This is a fairly complex problem which can easily be divide into several sub-problems. First and foremost it is necessary to find the means that will allow us to predict the development of the construction industry. It is possible to apply a production function, when an independent variable (production factory) is known and we look for a dependent variable i.e. the volume of production. The production functions for the coming years - if we know the volume of production factors - may indicate the future development of the industry. The subject of this paper is to find an answer to the first of these partial problems. The purpose of this is to be able to predict the development of the economy based on the development of the construction industry and to find the means by which these predictions can be made specifically based on production functions for the construction industry in the Czech Republic.


Key words: production factors, production funtions, construction industry, prediction.

## Introduction

Even in 2015 it can be said that the Czech economy continues to be in recession, a recession that has persisted for many years. Low GDP growth has been temporary and there are no signs of a long term trend towards improvement. Economists estimate the development of the economy primarily on the basis of macroeconomic variables such as GDP, balance of trade, inflation, unemployment, investment rate, etc.

Business and financial managers alike carefully observe such economic predictions. Their task is to financially manage specific enterprises in accordance with the objectives of their shareholders (according to the logic of increasing shareholder value). If we put forward the argument that every company tries to increase shareholder value (that is if they have no existential problems), this would mean that if the majority of enterprises in the economy are successful and are
meeting their basic objectives (i.e. they add value to the economy as a whole), it can be stated that the whole economy is successful and growing.

On the basis of this argument it would be logical to question whether it is then possible to predict the development of the economy according to the developments in, and forecasts from, the business environment. In other words, as the sum of the economic performances of business entities. The answer may be twofold. Firstly, we can assume that the development of the business environment heralds the development of the entire economy. On this basis we can predict the development of the entire economy on the results of business entities. Secondly, we can assume that higher entities behave with indifference to lower level entities.

In either case it is an extremely important moment for obtaining the relevant data and determining the production curve.

The aim of this paper is to determine the production curves for the construction industry in the Czech Republic and in its individual regions.

## 1. Material and methodology

The information given below about companies comes from the Albertina database. The data covers all the construction companies which operated on the Czech market between the years 2003-2011 and which fall into the classification CZNACE under Section F. The following activities are involved - construction of buildings, civil engineering and specialized construction activities.

The file contains a total of 67,492 rows of data in columns labelled:

- company registration no.;
- company name (15,189 companies);
- region
- list of annual financial statements between 2002-2011 (some companies did not submit financial statements in some years, some did it twice, three times or four times a year);
- sales (dependent variable y - production),
- $\quad$ sales costs + power consumption (explanatory variable $=$ regressor $x_{\mathbf{1}}-\mathbf{1}$ production factor);
- labour costs (explanatory variable $\boldsymbol{x}_{\mathbf{z}}-\mathbf{2}$ production factor),
- amortization of intangible and tangible assets (explanatory variable $x_{\mathbf{3}}-\mathbf{3}$ production factor);
- other operating expenses (explanatory variable $x_{\mathbf{L}}-\mathbf{4}$ production factor),
- interest (explanatory variable $x_{5}-5$ production factor).

The production value and production factors are stated in TCZK (thousands of Czech Crowns). These values can be zero or negative too. The data is completed with the following variables: dic (data identification number, $1-67492$, jv vector of ones),

$$
\begin{gather*}
\ln y=\ln (y+26481833), \ln x_{1}=\ln \left(x_{1}+27143333\right), \ln x_{2}= \\
=\ln \left(x_{2}+159923\right) \\
\ln x_{3}=\ln \left(x_{3}+65409\right)
\end{gather*}
$$

Linear models were derived by using the method of least squares (for this purpose the introduction of logarithmic transformations is required). The programs MS Office Excel, R, STATGRAPHICS and GiveWin2 were used. The data was loaded into an R program in the sentence: >LNPF <- read.table ("clipboard", header=TRUE, sep="\t", na.strings="NA", + dec=",", strip.white=TRUE).

Results - production function.
Based on the imported underlying data, basic data on variables was obtained in an $R$ program (year_k is a categorical variable year). The linear dependence of production on production factors is indicated by the correlation matrix.

## 2. Linear single-factor production function

The linear single-factor production function where $y$ is dependent on $x_{1}$ is subjected to an investigation into the shape of the regression line

$$
y_{i}=\beta_{0}+\beta_{1} x_{i}+\varepsilon i,
$$

where $\beta_{0}$ and $\beta_{1} \quad$ are the estimated regression parameters (the absolute term - the intercept - a regression coefficient describes the absolute flexibility of the production factor $y$ on $x_{1}$ )
$\varepsilon_{i}$ is a random (error component), $\mathrm{i}=1, \ldots, \mathrm{n}$ ( n is the number of measurements).
A regression line is determined on the basis of an estimation (offsetting) using the method of the least squares (hereinafter referred to as "LSM"). This method aims to minimize the sum of the squared residues i.e. the differences between the measured values and equivalent values measured on a straight line, using the software R.

The output is the following result:
$\operatorname{lm}\left(\right.$ formula $=y \sim x_{1}$, data $\left.=L N P F\right)$
The resulting values are shown in Table 1.
Table 1 - Results from the application of the LSM on production factor $\mathrm{x}_{1}$

|  | Estimate | Std. error | t -value | $\operatorname{Pr}(<\|\mathrm{t}\|)$ |  |
| :--- | :--- | :--- | :--- | :--- | :--- |
| Intercept | $3.19 e+03$ | $1.44 e+02$ | 22.05 | $<2 e-16$ | $* * *$ |
| $x_{1}$ | $1.216 e+00$ | $3.470 e-04$ | 3504.84 | $<2 e-16$ | $* * *$ |

Source: Author
Residual standard error: 37,400 on 67,490 degrees of freedom
Multiple R-squared: 0.9945 Adjusted R-squared: 0.9945
F-statistic: $1.228 \mathrm{e}+07$ on 1 and 67,490 DF, $p$-value: $<2.2 \mathrm{e}-16$
The estimated production function can be expressed as:
$y=3195+1.216 x_{1}+e$
where $e$ is a vector of residues (differences between measured and balanced values), which represent the estimates for random elements.

From small p-value (risks of an incorrect rejection test null hypothesis) significance tests it is evident that both the estimated regression parameters are statistically significant (production factor $x_{1}$ significantly acts linearly on
production function $y$ and at an average initial level of production of TCZK 3,195 when $\mathrm{x} \_1=0$ is significantly set). The slope can be interpreted as the absolute flexibility i.e. with every increase in production factor $x$ by a total of one thousand, output $y$ increases on average by TCZK 1,216 . The strength of dependence is expressed by the coefficient of determination $\mathrm{R}^{2}=0.9945$ This can be interpreted to mean that $99.45 \%$ of changes in production factor $x_{1}$ can be explained. This fact confirms the very strong linear dependence of production function y on factor $\boldsymbol{x}_{1}$. This is evident from the close proximity of the dots on the dot plot, which represent the measured values, to the balanced line. Other factors relating to changes in production function y can only explain a maximum of $0.55 \%$ of changes. This single-factor production function is therefore crucial. In mathematical statistics it has been shown that the estimates obtained through the LSM are impartial and consistent (approaching in probability the estimated parameters). In order to minimise selection errors so that other statistical induction (tests and confidence intervals) could be performed, the normality of the distribution of random components with zero value and a constant variance (i.e. so-called homoscedasticity of variances measurement $y$ ) and the linear independence of the individual measurement of output (i.e. between measurements is not a serial dependence - the so-called autocorrelation) are required. The verification of these conditions is usually done on residues graphically using residual analysis (or by statistical tests applied to the residues). Each program therefore counts residues and their standardized values (i.e. studentized residues, which in the case of normality should randomly fluctuate around zero, maximally within the range from -8 to +8 ). The dependence is usually shown by the studentized residues on the balanced values and serial numbers of measurement. If the residues do not move within that band (having the shape of open scissors), there is heteroscedasticity (opposite homoscedasticity) in the model. This can be dealt with by subjecting the estimate of the regression function to a so-called weighted LSM. If there is inertia in the residues there is an autocorrelation. In such cases estimates of the generalized LSM (e.g. in the model CORC- methods with offered econometric applications GiveWin2) are used instead of the ordinary LSM. Normality can be verified through a QQ-diagram (the dependency of empirical and theoretical quantiles of residues and normal quantiles lying on the line). If the points - showing the empirical quantiles - are close to that line, the residues are normally distributed. Graphs of residues are given in Figure 1, 2.


Figure 1 - Graphs of residues


Figure 1 - Graphs of residues
These graphs of studentized residues show that the normality of data is breached and in the model there is a certain level of heteroscedasticity in variance of residues (according to the previous dot plot, the level is not relatively high). These problems can be partially explained by the presence of relatively large amounts of low and high, outlying and extreme values in the data. This can also been seen in the last picture which is a box plot for the course of studentized residues. It is necessary - with discretion - to look at the possible use of other methods of statistical induction in the model and at the evaluation of the size of the standard errors of the estimated regression parameters. The static quality of the model could be improved by deleting these values, but the values of the estimated regression parameters would consequently change.

Further specification of the production function derived by year or by region is also of interest. If a qualitative (categorical) variable is to be included in the regression model, a so-called dummy variable must be used to determine its levels. These dummy variables have a value of 1 if the qualitative variable comes in at this level and otherwise the value 0 . In order to prevent multicollinearity in the regression model, the number of used dummy variables must be equal to the number of levels minus one. Any level at which a dummy variable has been omitted is called a reference variable (this is usually atthe first level). The estimated regression coefficients of dummy variables indicates how much the average value of the explanatory variable has changed while the level of dummy variables passed from the reference level to the intended level (e.g. the region has changed from JCK under consideration to another region, or the year changed from 2002 to another intended year). The model can include quantitative interactions with a qualitative variable (the estimated coefficient is interpreted as an amendment to the directive corresponding to the transition from the reference level to the intended level of the qualitative variable).

First, we assume the dependence on production factor $x_{1}$ and the year. The year is a categorical variable to levels in 2002, 2003, 2004, 2005, 2006, 2007, 2008, 2009, 2010 and 2011. Dummy variables $\mathrm{u}_{02}, \mathrm{u}_{03}, \ldots, \mathrm{u}_{11}$ can be introduced for
each year. The first dummy variable $\mathrm{u}_{02}$ is omitted. This variable for the year is considered to be level reference one. Other dummy variables $u_{03}$ to $u_{11}$ are included in the model and a production function predicted:

$$
\begin{gather*}
y=\beta_{0}+\beta_{1} x_{1}+\beta_{2} u_{03}+\beta_{3} u_{04}+\beta_{4} u_{05}+\cdots+\beta_{10} u_{11}+\beta_{11} x_{1} u_{03} \\
+\beta_{12} x_{1} u_{04}+\cdots+\beta_{19} x_{1} u_{11}+\varepsilon
\end{gather*}
$$

The equation can be established again in the program R1 using the LSM. The resulting parameters are listed in Table 2.

Table 2 - Resulting parameters of a production function for production factor $\mathrm{x}_{1}$ according to the individual years in the reference period

|  | Estimate | Std. error | t-value | $\operatorname{Pr}(<\|\mathrm{t}\|)$ |  |
| :--- | :--- | :--- | :--- | :--- | :--- |
| (Intercept) | $-3.160 \mathrm{e}+03$ | $6.406 \mathrm{e}+03$ | -0.493 | 0.622 |  |
| $x_{1}$ | $1.373 \mathrm{e}+00$ | $3.008 \mathrm{e}-03$ | 456.574 | $<2 \mathrm{e}-16$ | $* * *$ |
| year_k[T.2003] | $5.966 \mathrm{e}+03$ | $6.428 \mathrm{e}+03$ | 0.928 | 0.353 |  |
| year_k[T.2004] | $5.782 \mathrm{e}+03$ | $6.423 \mathrm{e}+03$ | 0.900 | 0.368 |  |
| year_k[T.2005] | $6.099 \mathrm{e}+03$ | $6.420 \mathrm{e}+03$ | 0.950 | 0.342 |  |
| year_k[T.2006] | $6.003 \mathrm{e}+03$ | $6.418 \mathrm{e}+03$ | 0.935 | 0.350 |  |
| year_k[T.2007] | $6.003 \mathrm{e}+03$ | $6.416 \mathrm{e}+03$ | 1.075 | 0.282 |  |
| year_k[T.2008] | $7.668 \mathrm{e}+03$ | $6.416 \mathrm{e}+03$ | 1.195 | 0.232 |  |
| year_k[T.2009] | $6.442 \mathrm{e}+03$ | $6.416 \mathrm{e}+03$ | 1.004 | 0.315 |  |
| year_k[T.2010] | $5.872 \mathrm{e}+03$ | $6.417 \mathrm{e}+03$ | 0.915 | 0.360 |  |
| year_k[T.2011] | $7.045 \mathrm{e}+03$ | $6.466 \mathrm{e}+03$ | 1.090 | 0.276 |  |
| $x_{1}:$ year_k[T.2003] | $-1.408 \mathrm{e}-01$ | $3.332 \mathrm{e}-03$ | -42.256 | $<2 \mathrm{e}-16$ | $* * *$ |
| $x_{1}:$ year_k[T.2004] | $-1.530 \mathrm{e}-01$ | $3.223 \mathrm{e}-03$ | -47.491 | $<2 \mathrm{e}-16$ | $* * *$ |
| $x_{1}:$ year_k[T.2005] | $-1.547 \mathrm{e}-01$ | $3.151 \mathrm{e}-03$ | -49.086 | $<2 \mathrm{e}-16$ | $* * *$ |
| $x_{1}:$ year_k[T.2006] | $-1.635 \mathrm{e}-01$ | $3.147 \mathrm{e}-03$ | -51.968 | $<2 \mathrm{e}-16$ | $* * *$ |
| $x_{1}:$ year_k[T.2007] | $-1.766 \mathrm{e}-01$ | $3.153 \mathrm{e}-03$ | -56.006 | $<2 \mathrm{e}-16$ | $* * *$ |
| $x_{1}:$ year_k[T.2008] | $-1.900 \mathrm{e}-01$ | $3.136 \mathrm{e}-03$ | -60.577 | $<2 \mathrm{e}-16$ | $* * *$ |
| $x_{1}:$ year_k[T.2009] | $-1.448 \mathrm{e}-01$ | $3.133 \mathrm{e}-03$ | -46.225 | $<2 \mathrm{e}-16$ | $* * *$ |
| $x_{1}:$ year_k[T.2010] | $-1.387 \mathrm{e}-01$ | $3.175 \mathrm{e}-03$ | -43.687 | $<2 \mathrm{e}-16$ | $* * *$ |
| $x_{1}:$ year_k[T.2011] | $-1.502 \mathrm{e}-01$ | $3.251 \mathrm{e}-03$ | $<2 \mathrm{e}-16$ | $<2 \mathrm{e}-16$ | $* * *$ |

Source: Author
Residual standard error: 36,000 on 67,472 degrees of freedom
Multiple R-squared: 0.9949, Adjusted R-squared: 0.9949
F-statistic: $6.983 \mathrm{e}+05$ on 19 and $67,472 \mathrm{DF}, \mathrm{p}$-value: $<2.2 \mathrm{e}-16$

[^0]Here, for example, year_k[T.2003] indicates u03, and x1:year_k[T.2003] indicates x 1 u 03 . The estimated production function is:

$$
\begin{align*}
y=-3160+ & 1.373 x_{1}+5966 u_{03}+5782 u_{04}+\cdots \\
& +7045 u_{11}-0.141 x 1_{1} u 03_{03}-0.153 x_{1} u_{04}-\cdots-0.1 x_{1} u_{1} \\
& +e
\end{align*}
$$

The production function for 2002 (reference period) can be determined if we substitute all the dummy variables for 0 . The production function for 2002 is therefore as follows:

$$
\begin{equation*}
y=-3160+1.373+e \tag{5}
\end{equation*}
$$

The production function for 2003 can be obtained by substituting $u 03=1$, and by substituting the other dummy variables for 0 . The production function for 2003 is therefore as follows:

$$
\begin{aligned}
& y=-3160+1.373 x_{1}+5966+0+\cdots+0-0.141 x_{1}-0-\cdots .-0+e=(6 \\
& 2806+\quad+1.232 x_{1}+e
\end{aligned}
$$

By continuing this process (i.e. by adding the coefficient of the respective dummy variable to the intercept, and by adding the coefficient ux 1 mutipled by the dummy variable to the slope) the equations for the production functions for years are as follows:

$$
\begin{align*}
& \text { 2004: } y=-3160+5782+1.373 x_{1}-0.153 x_{1}+e=2622+1.220 x_{1}+e \\
& 2005: y=-3160+6099+1.373 x_{1}-0.153 x_{1}+e=2939+1.218 x_{1}+e \\
& 2006: y=-3160+6003+1.373 x_{1}-0.164 x_{1}+e=2846+1.209 x_{1}+e \\
& 2007: y=-3160+6898+1.373 x_{1}-0.177 x_{1}+e=3739+1.196 x_{1}+e  \tag{7}\\
& 2008: y=-3160+7668+1.373 x_{1}-0.190 x_{1}+e=4508+1.186 x_{1}+e \\
& 2009: y=-3160+5782+1.373 x_{1}-0.145 x_{1}+e=3282+1.228 x_{1}+e \\
& 2010 ; y=-3160+5782+1.373 x_{1}-0.139 x_{1}+e=2712+1.234 x_{1}+e \\
& 2011: y=-3160+7045+1.373 x_{1}-0.150 x_{1}+e=3885+1.223 x_{1}+e
\end{align*}
$$

In summary it can be stated that the average levels of production (at $x_{1}=0$ ) are not statistically significant in the individual years (in 2002 there were only 33 highly variable figures). On the contrary, absolute flexibilities are statistically significant (marked by asterisks) and vary only slightly.

The R-square (coefficient of determination) rose to 0.9949 , i.e. $99.49 \%$ of all changes to production $y$ in the individual years is explained by factor $x_{1}$. The conclusions, subject to the verification of conditions, are similar to the previous model.

In the following part, linearsingle-factor production functions will be compared for production factor $x_{1}$ with the variable region.

Region JCK (South Bohemia) represents the basic level (alphabetically it is in the first place). Apart from the variable $x_{l}$, the explanatory dummy variables for the other regions ( $\left.u_{J C K}, u_{J M K}, u_{K V K}, u_{K V}, u_{K K}, u_{L K}, u_{M K}, u_{O K}, u_{P K}, u_{P K}, u_{S K}, u_{U K}, u_{Z K}, u_{P H A}\right)$
will be inserted into the model (designated in R as region[T.JMK], region[T.KK], ..., region[T.ZK]).

The results of the analysis are presented in Table 3.

Table 3 - Resulting parameters of production function for production factor x 1 in the individual regions

|  | Estimate | Std. error | t -value | $\operatorname{Pr}(>\|\mathrm{t}\|)$ |  |
| :--- | :--- | :--- | :--- | :--- | :--- |
| (Intercept) | $4.739 \mathrm{e}+03$ | $5.778 \mathrm{e}+02$ | 8.201 | $2.42 \mathrm{e}-16$ | $* * *$ |
| $x_{1}$ | $1.169 \mathrm{e}+00$ | $2.972 \mathrm{e}-03$ | 393.282 | $<2 \mathrm{e}-16$ | $* * *$ |
| region[T.JMK] | $-1.365 \mathrm{e}+03$ | $6.971 \mathrm{e}+02$ | -1.958 | 0.05028 |  |
| region[T.KK] | $-2.422 \mathrm{e}+03$ | $8.621 \mathrm{e}+02$ | -2.809 | 0.00497 | $* *$ |
| region[T.KV] | $-5.905 \mathrm{e}+03$ | $1.004 \mathrm{e}+03$ | -5.879 | $4.14 \mathrm{e}-09$ | $* * *$ |
| region[T.KVK] | $-1.647 \mathrm{e}+03$ | $1.191 \mathrm{e}+03$ | -1.383 | 0.16670 |  |
| region[T.LK] | $-1.289 \mathrm{e}+03$ | $9.089 \mathrm{e}+02$ | -1.418 | 0.15620 |  |
| region[T.MK] | $-3.432 \mathrm{e}+03$ | $7.382 \mathrm{e}+02$ | -4.650 | $3.33 \mathrm{e}-06$ | $* * *$ |
| region[T.OK] | $-5.484 \mathrm{e}+03$ | $9.453 \mathrm{e}+02$ | -5.801 | $6.61 \mathrm{e}-09$ | $* * *$ |
| region[T.PHA] | $-1.266 \mathrm{e}+03$ | $6.426 \mathrm{e}+02$ | -1.970 | 0.04881 | $*$ |
| region[T.PK] | $-4.222 \mathrm{e}+03$ | $7.562 \mathrm{e}+02$ | -5.583 | $2.37 \mathrm{e}-08$ | $* * *$ |
| region[T.SK] | $6.726 \mathrm{e}+02$ | $7.633 \mathrm{e}+02$ | 0.881 | 0.37823 |  |
| region[T.UK] | $-9.768 \mathrm{e}+02$ | $7.817 \mathrm{e}+02$ | -1.249 | 0.21150 |  |
| region[T.ZK] | $-3.605 \mathrm{e}+03$ | $8.547 \mathrm{e}+02$ | -4.218 | $2.47 \mathrm{e}-05$ | $* * *$ |
| $x_{1}:$ region[T.JMK] | $1.026 \mathrm{e}-02$ | $3.195 \mathrm{e}-03$ | 3.211 | 0.00132 | $* *$ |
| $x_{1}:$ :egion[T.KK] | $2.587 \mathrm{e}-02$ | $8.474 \mathrm{e}-03$ | 3.053 | 0.00227 | $* *$ |
| $x_{1}:$ region[T.KV] | $2.132 \mathrm{e}-01$ | $8.621 \mathrm{e}-03$ | 24.730 | $<2 \mathrm{e}-16$ | $* * *$ |
| $x_{1}:$ region[T.KVK] | $4.689 \mathrm{e}-02$ | $1.533 \mathrm{e}-02$ | 3.059 | 0.00222 | $* *$ |
| $x_{1}:$ region[T.LK] | $5.164 \mathrm{e}-03$ | $5.198 \mathrm{e}-03$ | 0.993 | 0.32051 |  |
| $x_{1}:$ region[T.MK] | $1.146 \mathrm{e}-01$ | $3.467 \mathrm{e}-03$ | 33.062 | $<2 \mathrm{e}-16$ | $* * *$ |
| $x_{1}:$ region[T.OK] | $1.529 \mathrm{e}-01$ | $6.404 \mathrm{e}-03$ | 23.883 | $<2 \mathrm{e}-16$ | $* * *$ |
| $x_{1}:$ region[T.PHA] | $4.890 \mathrm{e}-02$ | $2.995 \mathrm{e}-03$ | 16.327 | $<2 \mathrm{e}-16$ | $* * *$ |
| $x_{1}:$ region[T.PK] | $1.550 \mathrm{e}-01$ | $4.939 \mathrm{e}-03$ | 31.386 | $<2 \mathrm{e}-16$ | $* * *$ |
| $x_{1}:$ region[T.SK] | $-1.815 \mathrm{e}-02$ | $3.653 \mathrm{e}-03$ | -4.968 | $6.78 \mathrm{e}-07$ | $* * *$ |
| $x_{1}:$ region[T.UK] | $5.045 \mathrm{e}-02$ | $4.249 \mathrm{e}-03$ | 11.872 | $<2 \mathrm{e}-16$ | $* * *$ |
| $x_{1}:$ region[T.ZK] | $1.016 \mathrm{e}-01$ | $5.441 \mathrm{e}-03$ | 18.680 | $<2 \mathrm{e}-16$ | $* * *$ |

Source: Author
Residual standard error: 35,980 on 67,466 degrees of freedom.
Multiple R-squared: 0.9949
Adjusted R-squared: 0.9949
F-statistic: $5.31 \mathrm{e}+05$ on 25 and $67466 \mathrm{DF}, \mathrm{p}$-value: $<2.2 \mathrm{e}-16$
The production function for region JCK is:

$$
\begin{equation*}
y=4739+1.169 x+e \tag{8}
\end{equation*}
$$

The coefficients next to the dummy variables indicate how much the initial average production changes during the transition from region JCK to another region (it mostly declines, with the most significant decrease being TCZK 5,905
when moving from region JCK to region KV). However, what is more important are the changes to absolute flexibility that occur with regards to the product variables (the last twelve lines). For all the regions, except for region SK, the absolute flexibility is higher than in region JCK. Compared to region JCK, the highest increase (by TCZK 213) is represented by region KV. An overview of the resulting parameters in TCZK for the production function $y=b_{0}+b_{1} x_{1}$ is given in Table 4. From this overvew it is clear that most of the regression parameters are statistically significant.

Table4 - Overview of resulting parameters in TCZK for theproduction function

| $y=b_{0}+b_{1} x_{1}$ |  |  |  |  |
| :--- | :--- | :--- | :--- | ---: |
| region | change $b_{0}$ | $b_{0}$ | change $b_{1}$ | $b_{1}$ |
| JCK | 0 | 4739 | 0 | 1.169 |
| JMK | -1365 | 3374 | 0.010 | 1.179 |
| KK | -2422 | 2317 | 0.026 | 1.195 |
| KV | -5905 | -1166 | 0.213 | 1.382 |
| KVK | -1647 | 3092 | 0.047 | 1.216 |
| LK | -1289 | 3450 | 0.005 | 1.174 |
| MK | -3432 | 1307 | 0.115 | 1.284 |
| OK | -5484 | -745 | 0.152 | 1.321 |
| PHA | -1266 | 3473 | 0.049 | 1.218 |
| PK | -4222 | 517 | 0.155 | 1.324 |
| SK | 673 | 5412 | -0.018 | 1.151 |
| UK | -977 | 3762 | 0.050 | 1.219 |
| ZK | -3605 | 1134 | 0.102 | 1.271 |

Source: Author
The results of the conditions verification for the use of this model are analogous to those of the first model.

A brief description follows of the results of the linear single-factor model for production $y$ dependence on the other production factors. The comments on the results are analogous to the previous models.

## 3. Linear single-factor model for production $y$ dependence on production factor $x_{2}$

The result of the LSM application to production factor $x_{2}$ is given in Table 5.
Table 5 - Results of the LSM application to production factor $\mathrm{x}_{2}$

|  | Estimate | Std. error | t -value | $\operatorname{Pr}(>\|\mathrm{t}\|)$ |  |
| :--- | :--- | :--- | :--- | :--- | :--- |
| (Intercept) | 12087.12 | 1097.00 | 11.02 | $<2 \mathrm{e}-16$ | $* * *$ |
| $x_{2}$ | 11.55 | 0.03 | 384.90 | $<2 \mathrm{e}-1$ | $* * *$ |

Source: Author
Residual standard error: 283,100 on 67,490 degrees of freedom Multiple R-squared: 0.687 Adjusted R-squared: 0.687
F-statistic: $1.482 \mathrm{e}+05$ on 1 and 67490 DF, p-value: $<2.2 \mathrm{e}-16$
The single-factor model equation can be estimated again in the program R using the LSM. The resulting parameters are presented in Table 6.

Table 6 - Resulting parameters of the production function for production factor $\mathrm{x}_{2}$ according to the individual years in the reference period

|  | Estimate | Std. error | $\mathrm{t}-\mathrm{value}$ | $\operatorname{Pr}(<\|\mathrm{t}\|)$ |  |
| :--- | :--- | :--- | :--- | :--- | :--- |
| $($ Intercept $)$ | $7.600 \mathrm{e}+05$ | $5.354 \mathrm{e}+04$ | 14.195 | $<2 \mathrm{e}-16$ | $* * *$ |
| $x_{2}$ | $-1.634 \mathrm{e}+01$ | $3.162 \mathrm{e}+00$ | -5.166 | $2.39<2 \mathrm{e}-07$ | $* * *$ |
| year_k[T.2003] | $-7.138 \mathrm{e}+05$ | $5.367 \mathrm{e}+04$ | -13.300 | $<2 \mathrm{e}-16$ | $* * *$ |
| year_k[T.2004] | $-7.534 \mathrm{e}+05$ | $5.364 \mathrm{e}+04$ | -14.046 | $<2 \mathrm{e}-16$ | $* * *$ |
| year_k[T.2005] | $-7.398 \mathrm{e}+05$ | $5.362 \mathrm{e}+04$ | -13.797 | $<2 \mathrm{e}-16$ | $* * *$ |
| year_k[T.2006] | $7.581 \mathrm{e}+05$ | $5.361 \mathrm{e}+04$ | -14.140 | $<2 \mathrm{e}-16$ | $* * *$ |
| year_k[T.2007] | $-7.560 \mathrm{e}+05$ | $5.360 \mathrm{e}+04$ | -14.104 | $<2 \mathrm{e}-16$ | $* * *$ |
| year_k[T.2008] | $-7.576 \mathrm{e}+05$ | $5.360 \mathrm{e}+04$ | -14.136 | $<2 \mathrm{e}-16$ | $* * *$ |
| year_k[T.2009] | $-7.589 \mathrm{e}+05$ | $5.360 \mathrm{e}+04$ | -14.160 | $<2 \mathrm{e}-16$ | $* * *$ |
| year_k[T.2010] | $-7.586 \mathrm{e}+05$ | $5.360 \mathrm{e}+04$ | -14.151 | $<2 \mathrm{e}-16$ | $* * *$ |
| year_k[T.2011] | $-7.608 \mathrm{e}+05$ | $5.389 \mathrm{e}+04$ | -14.117 | $<2 \mathrm{e}-16$ | $* * *$ |
| $x_{2}:$ year_k[T.2003] | $2.111 \mathrm{e}+01$ | $3.162 \mathrm{e}+00$ | 6.6742 | $<2.51 \mathrm{e}-11$ | $* * *$ |
| $x_{2}:$ year_k[T.2004] | $3.033 \mathrm{e}+01$ | $3.164 \mathrm{e}+00$ | 9.586 | $<2 \mathrm{e}-16$ | $* * *$ |
| $x_{2}:$ year_k[T.2005] | $3.001 \mathrm{e}+01$ | $3.165 \mathrm{e}+00$ | 9.481 | $<2 \mathrm{e}-16$ | $* * *$ |
| $x_{2}:$ year_k[T.2006] | $3.149 \mathrm{e}+01$ | $3.163 \mathrm{e}+00$ | 9.956 | $<2 \mathrm{e}-16$ | $* * *$ |
| $x_{2}:$ year_k[T.2007] | $3.054 \mathrm{e}+01$ | $3.163 \mathrm{e}+00$ | 9.656 | $<2 \mathrm{e}-16$ | $* * *$ |
| $x_{2}:$ year_k[T.2008] | $2.977 \mathrm{e}+01$ | $3.163 \mathrm{e}+00$ | 9.412 | $<2 \mathrm{e}-16$ | $* * *$ |
| $x_{2}:$ year_k[T.2009] | $2.847 \mathrm{e}+01$ | $3.162 \mathrm{e}+00$ | 9.002 | $<2 \mathrm{e}-16$ | $* * *$ |
| $x_{2}:$ year_k[T.2010] | $2.788 \mathrm{e}+01$ | $3.162 \mathrm{e}+00$ | 8.817 | $<2 \mathrm{e}-16$ | $* * *$ |
| $x_{2}:$ year_k[T.2011] | $2.903 \mathrm{e}+01$ | $3.163 \mathrm{e}+00$ | 9.179 | $<2 \mathrm{e}-16$ | $* * *$ |

Source: Author
Residual standard error: 254,400 on 67,472 degrees of freedom
Multiple R-squared: 0.7472 Adjusted R-squared: 0.7471
F-statistic: $1.05 \mathrm{e}+04$ on 19 and 67472 DF , p-value: $<2.2 \mathrm{e}-16$
The results of the linear single-factor model for production $y$ dependence on production factor $x_{2}$ and the region are shown in Table 7.

Table7 - Resulting parameters of the production function for production factor $\mathrm{x}_{2}$ according to the individual regions

|  | Estimate | Std. error | t -value | $\operatorname{Pr}(>\|\mathrm{t}\|)$ |  |
| :--- | :--- | :--- | :--- | :--- | :--- |
| (Intercept) | 10028.5392 | 3969.9714 | 2.526 | 0.0115 | $*$ |


| $x_{2}$ | 9.7158 | 0.2067 | 47.011 | $<2 \mathrm{e}-16$ | $* * *$ |
| :--- | :--- | :--- | :--- | :--- | :--- |
| region[T.JMK] | -8169.2802 | 4794.1515 | -1.704 | 0.0884 |  |
| region[T.KK] | -5034.6940 | 5923.3758 | -0.850 | 0.3953 |  |
| region[T.KV] | -318.2433 | 6797.4757 | -0.047 | 0.9627 |  |
| region[T.KVK] | -163.2542 | 8185.7353 | -0.020 | 0.9841 |  |
| region[T.LK] | -2629.4228 | 6252.4120 | -0.421 | 0.6741 |  |
| region[T.MK] | 25197.9974 | 5052.9068 | 4.987 | $6.15 \mathrm{e}-07$ | $* * *$ |
| region[T.OK] | 11621.8814 | 6383.6814 | 1.821 | 0.0687 |  |
| region[T.PHA] | 7313.7758 | 4412.2246 | 1.658 | 0.0974 |  |
| region[T.PK] | -4971.1854 | 5193.9058 | -0.957 | 0.3385 |  |
| region[T.SK] | -6915.7810 | 5273.9969 | -1.311 | 0.1898 |  |
| region[T.UK] | -3053.9825 | 5371.1556 | -0.569 | 0.5696 |  |
| region[T.ZK] | 5887.2073 | 5872.6174 | 1.002 | 0.3161 |  |
| $x_{2}:$ region[T.JMK] | 5.8584 | 0.2441 | 24.001 | $<2 \mathrm{e}-16$ | $* * *$ |
| $x_{2}:$ region[T.KK] | 1.3965 | 0.6113 | 2.284 | 0.0224 | $*$ |
| $x_{2}:$ region[T.KV] | -0.4954 | 0.4454 | -1.112 | 0.2660 |  |
| $x_{2}:$ region[T.KVK] | -1.2031 | 0.9804 | -1.227 | 0.2197 |  |
| $x_{2}:$ region[T.LK] | 3.7006 | 0.4658 | 7.945 | $1.97 \mathrm{e}-15$ | $* * *$ |
| $x_{2}:$ region[T.MK] | -6.3060 | 0.2163 | -29.150 | $<2 \mathrm{e}-16$ | $* * *$ |
| $x_{2}:$ region[T.OK] | -3.1054 | 0.3047 | -10.191 | $<2 \mathrm{e}-16$ | $* * *$ |
| $x_{2}:$ region[T.PHA] | 3.6460 | 0.2090 | 17.446 | $<2 \mathrm{e}-16$ | $* * *$ |
| $x_{2}:$ region[T.PK] | 1.4353 | 0.3319 | 4.325 | $1.53 \mathrm{e}-05$ | $* * *$ |
| $x_{2}:$ region[T.SK] | 4.3173 | 0.3231 | 13.362 | $<2 \mathrm{e}-16$ | $* * *$ |
| $x_{2}:$ region[T.UK] | 1.4987 | 0.3009 | 4.981 | $6.35 \mathrm{e}-07$ | $* * *$ |
| $x_{2}:$ region[T.ZK] | -0.3238 | 0.4061 | -0.797 | 0.4253 |  |

Source: Author
Residual standard error: 246,600 on 67,466 degrees of freedom
Multiple R-squared: 0.7626 Adjusted R-squared: 0.7625
F-statistic: 8670 on 25 and 67466 DF, p-value: $<2.2 \mathrm{e}-16$

## 4. Linear single-factor model for production $\boldsymbol{y}$ dependence on production factor $x_{3}$

The results of the application of the LSM to production factor $x_{3}$ are given in Table 8.

Table 8 - Results of the application of the LSM to production factor $\mathrm{x}_{3}$

|  | Estimate | Std. error | t -value | $\operatorname{Pr}(>\|\mathrm{t}\|)$ |  |
| :--- | :--- | :--- | :--- | :--- | :--- |
| (Intercept) | 15664.223 | 1185.551 | 13.21 | $<2 \mathrm{e}-16$ | $* * *$ |
| $x_{3}$ | 46.855 | 0.137 | 342.01 | $<2 \mathrm{e}-16$ | $* * *$ |

Source: Author
Residual standard error: 306,000 on 67,490 degrees of freedom
Multiple R-squared: $0.6341 \quad$ Adjusted R-squared: 0.6341
F-statistic: $1.17 \mathrm{e}+05$ on 1 and 67490 DF, $p$-value: $<2.2 \mathrm{e}-16$

The single-factor model equation is estimated again in the program R using the LSM. The resulting parameters are presented in Table 9.
Table 9 - Resulting parameters of the production function for production factor $\mathrm{x}_{3}$ according to the individual years in the reference period

|  | Estimate | Std. error | t -value | $\operatorname{Pr}(<\|\mathrm{t}\|)$ |  |
| :--- | :--- | :--- | :--- | :--- | :--- |
| (Intercept) | $2.690 \mathrm{e}+04$ | $5.283 \mathrm{e}+04$ | 0.509 | 0.610638 |  |
| $x_{3}$ | $4.758 \mathrm{e}+01$ | $8.615 \mathrm{e}-01$ | 55.228 | $<2 \mathrm{e}-16$ | $* * *$ |
| year_k[T.2003] | $-1.634 \mathrm{e}+04$ | $5.302 \mathrm{e}+04$ | -0.308 | 0.757926 |  |
| year_k[T.2004] | $-1.795 \mathrm{e}+04$ | $5.297 \mathrm{e}+04$ | -0.339 | 0.734779 |  |
| year_k[T.2005] | $-1.971 \mathrm{e}+04$ | $5.295 \mathrm{e}+04$ | -0.372 | 0.709761 |  |
| year_k[T.2006] | $-1.795 \mathrm{e}+04$ | $5.293 \mathrm{e}+04$ | -0.339 | 0.734528 |  |
| year_k[T.2007] | $-4.710 \mathrm{e}+03$ | $5.292 \mathrm{e}+04$ | -0.089 | 0.929079 |  |
| year_k[T.2008] | $-2.688 \mathrm{e}+03$ | $5.291 \mathrm{e}+04$ | -0.051 | 0.959493 |  |
| year_k[T.2009] | $-1.208 \mathrm{e}+04$ | $5.291 \mathrm{e}+04$ | -0.228 | 0.819355 |  |
| year_k[T.2010] | $-1.282 \mathrm{e}+04$ | $5.292 \mathrm{e}+04$ | -0.242 | 0.808608 |  |
| year_k[T.2011] | $-2.047 \mathrm{e}+04$ | $5.333 \mathrm{e}+04$ | -0.384 | 0.701012 |  |
| $x_{3}:$ year_k[T.2003] | $1.154 \mathrm{e}+01$ | $1.083 \mathrm{e}+00$ | 10.657 | $<2 \mathrm{e}-16$ | $* * *$ |
| $x_{3}:$ year_k[T.2004] | $1.142 \mathrm{e}+01$ | $1.021 \mathrm{e}+00$ | 11.189 | $<2 \mathrm{e}-16$ | $* * *$ |
| $x_{3}:$ year_k[T.2005] | $9.608 \mathrm{e}+00$ | $1.021 \mathrm{e}+00$ | 11.189 | $<2 \mathrm{e}-16$ | $* * *$ |
| $x_{3}:$ year_k[T.2006] | $1.242 \mathrm{e}+01$ | $9.691 \mathrm{e}-01$ | 12.818 | $<2 \mathrm{e}-16$ | $* * *$ |
| $x_{3}:$ year_k[T.2007] | $-2.720 \mathrm{e}+00$ | $9.462 \mathrm{e}-01$ | -2.875 | 0.004040 | $* *$ |
| $x_{3}:$ year_k[T.2008] | $-7.875 \mathrm{e}+00$ | $9.270 \mathrm{e}-01$ | -8.495 | $<2 \mathrm{e}-16$ | $* * *$ |
| $x_{3}:$ year_k[T.2009] | $-7.197 \mathrm{e}+00$ | $9.154 \mathrm{e}-01$ | -7.862 | $3.85 \mathrm{e}-15$ | $* * *$ |
| $x_{3}:$ year_k[T.2010] | $-9.389 \mathrm{e}+00$ | $9.280 \mathrm{e}-01$ | -10.118 | $<2 \mathrm{e}-16$ | $* * *$ |
| $x_{3}:$ year_k[T.2011] | $3.438 \mathrm{e}+00$ | $9.767 \mathrm{e}-01$ | 3.520 | 0.000431 | $* * *$ |

Source: Author
Residual standard error: 297,500 on 67,472 degrees of freedom
Multiple R-squared: 0.6544 Adjusted R-squared: 0.6543
F-statistic: 6724 on 19 and 67472 DF, p-value: $<2.2 \mathrm{e}-16$
The results of the linear single-factor model for production $y$ dependence on production factor $x_{3}$ and the region are given in Table 10.

Table 10 - Resulting parameters of the production function for production factor $\mathrm{x}_{3}$ according to the individual regions

|  | Estimate | Std. error | t -value | $\operatorname{Pr}(>\|\mathrm{t}\|)$ |  |
| :--- | :--- | :--- | :--- | :--- | :--- |
| (Intercept) | 12497.7201 | 4595.1744 | 2.720 | 0.00653 | $* *$ |
| $x_{3}$ | 45.1346 | 1.2388 | 36.433 | $<2 \mathrm{e}-16$ | $* * *$ |
| region[T.JMK] | 24710.4797 | 5535.5871 | 4.464 | $8.06 \mathrm{e}-06$ | $* * *$ |
| region[T.KK] | -2580.7511 | 6794.7052 | -0.380 | 0.70408 |  |
| region[T.KV] | 12672.1787 | 7746.6942 | 1.636 | 0.10188 |  |


| region[T.KVK] | 6186.4484 | 9465.0203 | 0.654 | 0.51336 |  |
| :--- | :--- | :--- | :--- | :--- | :--- |
| region[T.LK] | 19577.5231 | 7174.8805 | 2.729 | 0.00636 | $* *$ |
| region[T.MK] | 12562.4482 | 5848.4599 | 2.148 | 0.03172 | $*$ |
| region[T.OK] | -3255.9070 | 7528.1553 | -0.432 | 0.66538 |  |
| region[T.PHA] | 7267.2783 | 5105.1992 | 1.424 | 0.15459 |  |
| region[T.PK] | 9521.9270 | 5959.9292 | 1.598 | 0.11012 |  |
| region[T.SK] | -3366.2787 | 6067.6280 | -0.555 | 0.57904 |  |
| region[T.UK] | 13776.5131 | 6184.6876 | 2.228 | 0.02592 | $*$ |
| region[T.ZK] | 15453.6051 | 6686.1185 | 2.311 | 0.02082 | $*$ |
| $x_{3}:$ region[T.JMK] | -22.7433 | 1.3178 | -17.258 | $<2 \mathrm{e}-16$ | $* * *$ |
| $x_{3}:$ region[T.KK] | 0.1404 | 3.0459 | 0.046 | 0.96325 |  |
| $x_{3}:$ region[T.KV] | -18.5642 | 2.0231 | -9.176 | $<2 \mathrm{e}-16$ | $* * *$ |
| $x_{3}:$ region[T.KVK] | -12.8074 | 6.9697 | -1.838 | 0.06613 |  |
| $x_{3}:$ region[T.LK] | -33.8729 | 1.6180 | -20.936 | $<2 \mathrm{e}-16$ | $* * *$ |
| $x_{3}:$ region[T.MK] | -16.5264 | 1.3171 | -12.548 | $<2 \mathrm{e}-16$ | $* * *$ |
| $x_{3}:$ region[T.OK] | 23.0879 | 3.3662 | 6.859 | $7.00 \mathrm{e}-12$ | $* * *$ |
| $x_{3}:$ region[T.PHA] | 9.0754 | 1.2479 | 7.272 | $3.57 \mathrm{e}-13$ | $* * *$ |
| $x_{3}:$ region[T.PK] | -15.7783 | 1.5823 | -9.972 | $<2 \mathrm{e}-16$ | $* * *$ |
| $x_{3}:$ region[T.SK] | 11.5434 | 1.6022 | 7.205 | $5.87 \mathrm{e}-13$ | $* * *$ |
| $x_{3}:$ :region[T.UK] | -17.5910 | 1.5401 | -11.422 | $<2 \mathrm{e}-16$ | $* * *$ |
| $x_{3}:$ region[T.ZK] | -23.7210 | 1.4542 | -16.312 | $<2 \mathrm{e}-16$ | $* * *$ |

Source: Author
Residual standard error: 284,800 on 67,466 degrees of freedom
Multiple R-squared: $0.6832 \quad$ Adjusted R-squared: 0.6831
F-statistic: 5820 on 25 and 67466 DF, p-value: $<2.2 \mathrm{e}-16$

## 5. Linear single-factor model for production $\boldsymbol{y}$ dependence on production factor $\boldsymbol{x}_{4}$

The results of the application of the LSM to production factor $x_{4}$ are given in Table 11.

Table 11 - The results of the application of the LSM to production factor $\mathrm{x}_{4}$

|  | Estimate | Std. error | t -value | $\operatorname{Pr}(>\|\mathrm{t}\|)$ |  |
| :--- | :--- | :--- | :--- | :--- | :--- |
| (Intercept) | $4.121 \mathrm{e}+04$ | $1.693 \mathrm{e}+03$ | 24.35 | $<2 \mathrm{e}-16$ | $* * *$ |
| $\mathrm{X}_{4}$ | $1.654 \mathrm{e}+01$ | $1.105 \mathrm{e}-01$ | 149.67 | $<2 \mathrm{e}-16$ | $* * *$ |

Source: Author
Residual standard error: 438,400 on 67,490 degrees of freedom
Multiple R-squared: 0.2492 Adjusted R-squared: 0.2492
F-statistic: $2.24 \mathrm{e}+04$ on 1 and 67490 DF, p-value: $<2.2 \mathrm{e}-16$
The single-factor model equation is estimated in the program R using the LSM. The resulting parameters are given in Table 12.

Table 12 - Resulting parameters of the production function for production factor $\mathrm{x}_{4}$ according to the individual years in the reference period

|  | Estimate | Std. error | $\mathrm{t}-\mathrm{value}$ | $\operatorname{Pr}(\langle \| \mathrm{t} \mid)$ |  |
| :--- | :--- | :--- | :--- | :--- | :--- |
| (Intercept) | -466066.99 | 78477.25 | -5.939 | $2.88 \mathrm{e}-09$ | $* * *$ |
| $x_{4}$ | 424.43 | 11.78 | 36.023 | $<2 \mathrm{e}-16$ | $* * *$ |
| year_k[T.2003] | 523006.24 | 78718.49 | 6.644 | $3.08 \mathrm{e}-11$ | $* * *$ |
| year_k[T.2004] | 499575.35 | 78664.24 | 6.351 | $2.16 \mathrm{e}-10$ | $* * *$ |
| year_k[T.2005] | 501582.58 | 78631.73 | 6.379 | $1.80 \mathrm{e}-10$ | $* * *$ |
| year_k[T.2006] | 492918.73 | 78612.39 | 6.270 | $3.63 \mathrm{e}-10$ | $* * *$ |
| year_k[T.2007] | 508932.28 | 78612.39 | 6.476 | $9.51 \mathrm{e}-11$ | $* * *$ |
| year_k[T.2008] | 512025.04 | 78587.12 | 6.515 | $7.30 \mathrm{e}-11$ | $* * *$ |
| year_k[T.2009] | 507968.62 | 78583.01 | 6.464 | $1.03 \mathrm{e}-10$ | $* * *$ |
| year_k[T.2010] | 501930.24 | 78597.36 | 6.386 | $1.71 \mathrm{e}-10$ | $* * *$ |
| year_k[T.2011] | 478717.01 | 79133.67 | 6.049 | $1.46 \mathrm{e}-09$ | $* * *$ |
| $x_{4}:$ year_k[T.2003] | -414.23 | 11.79 | -35.140 | $<2 \mathrm{e}-16$ | $* * *$ |
| $x_{4}:$ year_k[T.2004] | -414.23 | 11.79 | -33.779 | $<2 \mathrm{e}-16$ | $* * *$ |
| $x_{4}:$ year_k[T.2005] | -398.27 | 11.79 | -34.467 | $<2 \mathrm{e}-16$ | $* * *$ |
| $x_{4}:$ year_k[T.2006] | -406.20 | 11.79 | -32.850 | $<2 \mathrm{e}-16$ | $* * *$ |
| $x_{4}:$ year_k[T.2007] | -387.29 | 11.79 | -34.605 | $<2 \mathrm{e}-16$ | $* * *$ |
| $x_{4}:$ year_k[T.2008] | -410.87 | 11.79 | -34.862 | $<2 \mathrm{e}-16$ | $* * *$ |
| $x_{4}:$ year_k[T.2009] | -414.01 | 11.79 | -35.133 | $<2 \mathrm{e}-16$ | $* * *$ |
| $x_{4}:$ year_k[T.2010] | -410.13 | 11.79 | -34.799 | $<2 \mathrm{e}-16$ | $* * *$ |
| $x_{4}:$ year_k[T.2011] | -375.80 | 11.81 | -31.826 | $<2 \mathrm{e}-16$ | $* * *$ |

Source: Author
Residual standard error: 417,200 on 67,472 degrees of freedom
Multiple R-squared: 0.3202 Adjusted R-squared: 0.32
F-statistic: 1672 on 19 and 67472 DF, p-value: $<2.2 \mathrm{e}-16$
The results of the linear single-factor model for production $y$ dependence on production factor $x_{4}$ and the regions are given in Table 13.

Table 13 - The resulting parameters of the production function for production factor $\mathrm{x}_{4}$ according to the individual regions

|  | Estimate | Std. error | t -value | $\operatorname{Pr}(>\|\mathrm{t}\|)$ |  |
| :--- | :--- | :--- | :--- | :--- | :--- |
| (Intercept) | 30474.2785 | 6878.5812 | 4.430 | $9.42 \mathrm{e}-06$ | $* * *$ |
| $x_{4}$ | 22.0101 | 1.1595 | 18.983 | $<2 \mathrm{e}-16$ | $* * *$ |
| region[T.JMK] | 12549.1052 | 8306.1418 | 1.511 | 0.13084 |  |
| region[T.KK] | -7824.0283 | 10143.4349 | -0.771 | 0.44051 |  |
| region[T.KV] | 9197.5592 | 11534.7329 | 0.797 | 0.42523 |  |
| region[T.KVK] | -3201.0797 | 13666.7446 | -0.234 | 0.81481 |  |
| region[T.LK] | 3680.5467 | 10782.3806 | 0.341 | 0.73284 |  |


| region[T.MK] | 1629.7708 | 8798.4523 | 0.185 | 0.85305 |  |
| :--- | :--- | :--- | :--- | :--- | :--- |
| region[T.OK] | 1224.2068 | 11135.6782 | 0.110 | 0.91246 |  |
| region[T.PHA] | 30473.9353 | 7660.2720 | 3.978 | $6.95 \mathrm{e}-05$ | $* * *$ |
| region[T.PK] | 6020.2783 | 8923.0732 | 0.675 | 0.49988 |  |
| region[T.SK] | 8920.6082 | 9077.8329 | 0.983 | 0.32577 |  |
| region[T.UK] | 8550.1696 | 9281.7580 | 0.921 | 0.35696 |  |
| region[T.ZK] | 3099.7348 | 10063.4451 | 0.308 | 0.75807 |  |
| $x_{4}:$ region[T.JMK] | -10.8739 | 1.1924 | -9.119 | $<2 \mathrm{e}-16$ | $* * *$ |
| $x_{4}:$ region[T.KK] | -0.9454 | 3.4983 | -0.270 | 0.78696 |  |
| $x_{4}:$ region[T.KV] | -14.7233 | 1.4047 | -10.482 | $<2 \mathrm{e}-16$ | $* * *$ |
| $x_{4}:$ region[T.KVK] | -16.6689 | 1.8883 | -8.828 | $<2 \mathrm{e}-16$ | $* * *$ |
| $x_{4}:$ region[T.LK] | -15.1489 | 1.3772 | -11.000 | $<2 \mathrm{e}-16$ | $* * *$ |
| $x_{4}:$ region[T.MK] | -0.7168 | 1.3206 | -0.543 | 0.58726 |  |
| $x_{4}:$ region[T.OK] | 8.6624 | 3.0027 | 2.885 | 0.00392 | $* *$ |
| $x_{4}:$ region[T.PHA] | -2.4618 | 1.1677 | -2.108 | 0.03501 | $*$ |
| $x_{4}:$ region[T.PK] | -8.7781 | 1.5818 | -5.549 | $2.88 \mathrm{e}-08$ | $* * *$ |
| $x_{4}:$ region[T.SK] | -12.9655 | 1.2073 | -10.739 | $<2 \mathrm{e}-16$ | $* * *$ |
| $x_{4}:$ region[T.UK] | -8.0794 | 1.6312 | -4.953 | $7.33 \mathrm{e}-07$ | $* * *$ |
| $x_{4}:$ region[T.ZK] | -5.3507 | 1.5395 | -3.476 | 0.00051 | $* * *$ |

Source: Author
Residual standard error: 432,400 on 67,466 degrees of freedom
Multiple R-squared: 0.2699 Adjusted R-squared: 0.2697
F-statistic: 997.8 on 25 and 67466 DF, p-value: $<2.2 \mathrm{e}-16$

## 6. Linear single-factor model for production $\boldsymbol{y}$ dependence on production factor $x_{5}$

The results of the application of the LSM to production factor $x_{5}$ are given in Table 14.

Table 14 - The results of the application of the LSM to production factor $\mathrm{x}_{5}$

|  | Estimate | Std. error | t -value | $\operatorname{Pr}(>\|\mathrm{t}\|)$ |  |
| :--- | :--- | :--- | :--- | :--- | :--- |
| (Intercept) | $4.979 \mathrm{e}+04$ | $1.901 \mathrm{e}+03$ | 26.20 | $<2 \mathrm{e}-16$ | $* * *$ |
| $x_{5}$ | $4.689 \mathrm{e}+01$ | $7.396 \mathrm{e}-$ | 63.40 | $<2 \mathrm{e}-16$ | $* * *$ |

Source: Author
Residual standard error: 491,500 on 67,490 degrees of freedom
Multiple R-squared: 0.05621 Adjusted R-squared: 0.0562
F-statistic: 4020 on 1 and 67490 DF, $p$-value: $<2.2 \mathrm{e}-16$
The single-factor model equation is estimated in the program R using the LSM. The resulting parameters are given in Table 15.

Table 15 - The resulting parameters of the production function for production factor $\mathrm{x}_{5}$ according to the individual years in the reference period

|  | Estimate | Std. error | t-value | $\operatorname{Pr}(<\|t\|)$ |  |
| :--- | :--- | :--- | :--- | :--- | :--- |
| $($ Intercept $)$ | 729390.6 | 99265.6 | 7.348 | $2.04 \mathrm{e}-13$ | $* * *$ |
| $x_{5}$ | -581.5 | 241.4 | -2.409 | 0.015992 | $*$ |
| year_k[T.2003] | -670336.4 | 99527.0 | -6.735 | $1.65 \mathrm{e}-11$ | $* * *$ |
| year_k[T.2004] | -672406.0 | 99466.8 | -6.760 | $1.39 \mathrm{e}-11$ | $* * *$ |
| year_k[T.2005] | -671426.2 | 99431.8 | -6.753 | $1.46 \mathrm{e}-11$ | $* * *$ |
| year_k[T.2006] | -674466.9 | 99410.9 | -6.785 | $1.17 \mathrm{e}-11$ | $* * *$ |
| year_k[T.2007] | -682290.2 | 99389.5 | -6.865 | $6.72 \mathrm{e}-12$ | $* * *$ |
| year_k[T.2008] | -689796.7 | 99384.5 | -6.941 | $3.94 \mathrm{e}-12$ | $* * *$ |
| year_k[T.2009] | -687866.6 | 99379.9 | -6.922 | $4.51 \mathrm{e}-12$ | $* * *$ |
| year_k[T.2010] | -682673.9 | 99394.8 | -6.868 | $6.55 \mathrm{e}-12$ | $* * *$ |
| year_k[T.2011] | -711993.9 | 99978.3 | -7.121 | $1.08 \mathrm{e}-12$ | $* * *$ |
| $x_{5}:$ year_k[T.2003] | 650.3 | 241.4 | 2.693 | 0.007075 | $* *$ |
| $x_{5}:$ year_k[T.2004] | 620.3 | 241.4 | 2.570 | 0.010185 | $*$ |
| $x_{5}:$ year_k[T.2005] | 615.3 | 241.4 | 2.549 | 0.010807 | $*$ |
| $x_{5}:$ year_k[T.2006] | 620.0 | 241.4 | 2.568 | 0.010217 | $*$ |
| $x_{5}:$ year_k[T.2007] | 637.3 | 241.4 | 2.640 | 0.008293 | $* *$ |
| $x_{5}:$ year_k[T.2008] | 650.8 | 241.4 | 2.696 | 0.007016 | $* *$ |
| $x_{5}:$ year_k[T.2009] | 627.2 | 241.4 | 2.598 | 0.009370 | $* *$ |
| $x_{5}:$ year_k[T.2010] | 604.6 | 241.4 | 2.505 | 0.012265 | $*$ |
| $x_{5}:$ year_k[T.2011] | 888.9 | 241.5 | 3.680 | 0.000233 | $* * *$ |

Source: Author
Residual standard error: 486,400 on 67,472 degrees of freedom
Multiple R-squared: 0.07596 Adjusted R-squared: 0.0757
F-statistic: 291.9 on 19 and 67472 DF, p-value: $<2.2 \mathrm{e}-16$
The results of the linear single-factor model for production $y$ dependence on production factor $x_{5}$ and the regions are given in Table 16.

Table 16 - The resulting parameters of the production function for production factor $\mathrm{x}_{5}$ according to the individual regions

|  | Estimate | Std. error | t -value | $\operatorname{Pr}(>\|\mathrm{t}\|)$ |  |
| :--- | :--- | :--- | :--- | :--- | :--- |
| (Intercept) | 15315.416 | 7781.765 | 1.968 | 0.04906 | $*$ |
| $x_{5}$ | 213.256 | 12.011 | 17.756 | $<2 \mathrm{e}-16$ | $* * *$ |
| region[T.JMK] | 10699.029 | 9355.929 | 1.144 | 0.25281 |  |
| region[T.KK] | 2565.560 | 11365.226 | 0.226 | 0.82141 |  |
| region[T.KV] | 9700.561 | 13138.837 | 0.738 | 0.46033 |  |
| region[T.KVK] | 8093.617 | 15707.174 | 0.515 | 0.60636 |  |
| region[T.LK] | 6384.239 | 12077.346 | 0.529 | 0.59708 |  |


| region[T.MK] | 14345.605 | 9900.675 | 1.449 | 0.14736 |  |
| :--- | :--- | :--- | :--- | :--- | :--- |
| region[T.OK] | 7443.801 | 12512.836 | 0.595 | 0.55192 |  |
| region[T.PHA] | 61624.631 | 8639.664 | 7.133 | $9.94 \mathrm{e}-13$ | $* * *$ |
| region[T.PK] | 26911.970 | 9990.610 | 2.694 | 0.00707 | $* *$ |
| region[T.SK] | 10017.302 | 10203.453 | 0.982 | 0.32622 |  |
| region[T.UK] | 18136.686 | 10431.187 | 1.739 | 0.08209 |  |
| region[T.ZK] | 17687.625 | 11335.737 | 1.560 | 0.11868 |  |
| $x_{5}:$ region[T.JMK] | -25.694 | 12.694 | -2.024 | 0.04297 | $*$ |
| $x_{5}:$ region[T.KK] | -102.818 | 17.683 | -5.815 | $6.10 \mathrm{e}-09$ | $* * *$ |
| $x_{5}:$ region[T.KV] | -49.488 | 22.262 | -2.223 | 0.02622 | $*$ |
| $x_{5}:$ region[T.KVK] | -111.359 | 47.026 | -2.368 | 0.01789 | $*$ |
| $x_{5}:$ region[T.LK] | -132.696 | 12.864 | -10.315 | $<2 \mathrm{e}-16$ | $* * *$ |
| $x_{5}:$ region[T.MK] | -42.089 | 14.162 | -2.972 | 0.00296 | $* *$ |
| $x_{5}:$ region[T.OK] | 6.939 | 21.838 | 0.318 | 0.75066 |  |
| $x_{5}:$ region[T.PHA] | -162.319 | 12.051 | -13.469 | $<2 \mathrm{e}-16$ | $* * *$ |
| $x_{5}:$ region[T.PK] | -204.096 | 12.072 | -16.906 | $<2 \mathrm{e}-16$ | $* * *$ |
| $x_{5}:$ region[T.SK] | -74.997 | 12.766 | -5.874 | $4.26 \mathrm{e}-09$ | $* * *$ |
| $x_{5}:$ region[T.UK] | -82.646 | 15.458 | -5.346 | $9.00 \mathrm{e}-08$ | $* * *$ |
| $x_{5}:$ region[T.ZK] | -95.747 | 17.291 | -5.537 | $3.08 \mathrm{e}-08$ | $* * *$ |

Source: Author
Residual standard error: 479,500 on 67,466 degrees of freedom
Multiple R-squared: $0.1024 \quad$ Adjusted R-squared: 0.102
F-statistic: 307.8 on 25 and 67466 DF, p-value: $<2.2 \mathrm{e}-16$

## Conclusion

The objective of the paper was to determine the production curves for the construction industry in the Czech Republic and in its individual regions. The objective was achieved. The result is a production function for the construction industry at a regional and national level.

A tool is now at our disposal which allows us to predict the development of the construction industry in the Czech Republic in the years to come. By continuing to predict the values for the coming years we can subsequently compare them with the actual results in the field. In so doing, we can determine the degree of accuracy of the predictions for production and the development of the construction industry over time, both at regional and national levels.


[^0]:    ${ }^{I}$ Using the application R has the advantage that the necessary dummy variables will be introduced into the model automatically. It is not necessary to generate them.

